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Ethnicity and outcomes from COVID-19: the ISARIC CCP-UK prospective observational cohort study of hospitalised patients --Manuscript Draft--

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Abstract:	<p>Background Reports of ethnic inequalities in COVID-19 outcomes are conflicting and the reasons for any differences in outcomes are unclear. We investigated ethnic inequalities in critical care admission patterns, the need for invasive mechanical ventilation (IMV), and in-hospital mortality, among hospitalised patients with COVID-19.</p> <p>Methods We undertook a prospective cohort study in which dedicated research staff recruited hospitalised patients with suspected/confirmed COVID-19 from 260 hospitals across England, Scotland and Wales, collecting data directly and from records between 6th February and 8th May 2020 with follow-up until 22nd May 2020. Analysis used hierarchical regression models accounting for confounding, competing risks, and clustering of patients in hospitals. Potential mediators for death were explored with a three-way decomposition mediation analysis.</p> <p>Findings Of 34,986 patients enrolled, 30,693 (88%) had ethnicity recorded: South Asian (1,388, 5%), East Asian (266, 1%), Black (1,094, 4%), Other Ethnic Minority (2,398, 8%) (collectively Ethnic Minorities), and White groups (25,547, 83%). Ethnic Minorities were younger and more likely to have diabetes (type 1/type 2) but had fewer other comorbidities such as chronic heart disease or dementia than the White group. No difference was seen between ethnic groups in the time from symptom onset to hospital admission, nor in illness severity at admission. Critical care admission was more common in South Asian (odds ratio 1.28, 95% confidence interval 1.09 to 1.52), Black (1.36, 1.14 to 1.62), and Other Ethnic Minority (1.29, 1.13 to 1.47) groups compared to the White group, after adjusting for age, sex and location. This was broadly unchanged after adjustment for deprivation and comorbidities. Patterns were similar for IMV. Higher adjusted mortality was seen in the South Asian (hazard ratio 1.19, 1.05 to 1.36), but not East Asian (1.00, 0.74 to 1.35), Black (1.05, 0.91 to 1.26) or Other Ethnic Minority (0.99, 0.89 to 1.10) groups, compared to the White group. 18% (95% CI, 9% to 56%) of the excess mortality in South Asians was mediated by pre-existing diabetes.</p> <p>Interpretation Ethnic Minorities in hospital with COVID-19 were more likely to be admitted to critical care and receive IMV than Whites, despite similar disease severity on admission, similar duration of symptoms, and being younger with fewer comorbidities. South Asians are at greater risk of dying, due at least in part to a higher prevalence of pre-existing diabetes.</p> <p>Funding National Institute for Health Research and Medical Research Council.</p>

Ethnicity and outcomes from COVID-19: the ISARIC CCP-UK prospective observational cohort study of hospitalised patients

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Summary

Background

Reports of ethnic inequalities in COVID-19 outcomes are conflicting and the reasons for any differences in outcomes are unclear. We investigated ethnic inequalities in critical care admission patterns, the need for invasive mechanical ventilation (IMV), and in-hospital mortality, among hospitalised patients with COVID-19.

Methods

We undertook a prospective cohort study in which dedicated research staff recruited hospitalised patients with suspected/confirmed COVID-19 from 260 hospitals across England, Scotland and Wales, collecting data directly and from records between 6th February and 8th May 2020 with follow-up until 22nd May 2020. Analysis used hierarchical regression models accounting for confounding, competing risks, and clustering of patients in hospitals. Potential mediators for death were explored with a three-way decomposition mediation analysis.

Findings

Of 34,986 patients enrolled, 30,693 (88%) had ethnicity recorded: South Asian (1,388, 5%), East Asian (266, 1%), Black (1,094, 4%), Other Ethnic Minority (2,398, 8%) (collectively Ethnic Minorities), and White groups (25,547, 83%). Ethnic Minorities were younger and more likely to have diabetes (type 1/type 2) but had fewer other comorbidities such as chronic heart disease or dementia than the White group. No difference was seen between ethnic groups in the time from symptom onset to hospital admission, nor in illness severity at admission.

Critical care admission was more common in South Asian (odds ratio 1.28, 95% confidence interval 1.09 to 1.52), Black (1.36, 1.14 to 1.62), and Other Ethnic Minority (1.29, 1.13 to 1.47) groups compared to the White group, after adjusting for age, sex and location. This was broadly unchanged after adjustment for deprivation and comorbidities. Patterns were similar for IMV.

Higher adjusted mortality was seen in the South Asian (hazard ratio 1.19, 1.05 to 1.36), but not East Asian (1.00, 0.74 to 1.35), Black (1.05, 0.91 to 1.26) or Other Ethnic Minority (0.99, 0.89 to 1.10) groups, compared to the White group. 18% (95% CI, 9% to 56%) of the excess mortality in South Asians was mediated by pre-existing diabetes.

Interpretation

Ethnic Minorities in hospital with COVID-19 were more likely to be admitted to critical care and receive IMV than Whites, despite similar disease severity on admission, similar duration of symptoms, and being younger with fewer comorbidities. South Asians are at greater risk of dying, due at least in part to a higher prevalence of pre-existing diabetes.

Funding

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Research in Context

Evidence before this study:

Evidence is emerging of an association between ethnicity and increased mortality in patients with COVID-19, as shown in a rapid evidence review by McQuillan *et al.* In a cohort study on the relative risk of COVID-19 infection by ethnic group, Niedzwiedz *et al* linked UK Biobank data to show that Black (relative risk 4.01 [95% confidence interval 2.92 to 5.12]), South Asians (2.11 [1.43 to 3.10]) and White Irish (1.60 [1.08 to 2.38]) were more likely to have confirmed infection and were more likely to be hospitalised compared to White British. This relationship persisted after controlling for socioeconomic deprivation. In another study, de Noronha compared actual vs. expected hospital deaths and all ethnic groups other than White British and White Irish were found to have an age adjusted excess mortality. The UK Office of National Statistics used linked census and mortality records to show that age-adjusted COVID-19 -related death was more four times more likely in Black individuals (odds ratio 4.2 [3.81 to 4.63]). When accounting for socioeconomic differences, Black (1.93 [1.70 to 2.18]) and Bangladeshi/Pakistani (1.81 [1.55 to 2.11]) individuals were still at significantly increased risk of death compared to the White group. The Intensive Care National Audit & Research Centre (ICNARC) reported that 33% of patients critically ill with COVID-19 in England, Wales, and Northern Ireland are from non-White groups, compared to 12% with non-COVID-19 viral pneumonia in 2017-19.

Added value of this study

This is the largest prospective study of patients in hospital with COVID-19. We provide granular data on the stark differences in age and levels of comorbidity between White and Ethnic Minority people in hospital with COVID-19. We show high levels of use of critical care and invasive mechanical ventilation (IMV) in Ethnic Minorities compared with Whites, even after adjustment for age, sex, deprivation, and comorbidities, despite similar disease severity at presentation and similar times from symptom onset to hospital admission. We found a higher likelihood of death in hospital among South Asians with COVID-19, and for the first time show that a significant proportion of this risk was attributable to higher prevalence of diabetes.

Implications of all the available evidence

These findings have important implications for policy. Ethnicity should be accounted for in the prioritisation of prevention treatment and future vaccination policy. South Asians are over-represented in frontline key worker and public-interacting occupations, and guidance and policies should take this factor into account. Work is on-going to devise risk assessment tools to help mitigate the likelihood of coronavirus infection and manage the easing of lockdown restrictions. Careful consideration needs to be given in these tools to the weighting of ethnicity and mediators of risk such as diabetes. Further research integrating primary and secondary care data is key to improving the understanding of other drivers of poor outcomes among Ethnic Minorities.

Introduction

The novel severe acute respiratory syndrome coronavirus SARS-CoV-2 has caused a pandemic, with 5.8 million reported as infected and over 360,000 reported to have died across the world.¹ Studies from China, Italy and the UK have consistently reported that advanced age, comorbidity, and male sex are associated with increased mortality.²⁻⁴

COVID-19 deaths appear to disproportionately affect people from South Asian and Black ethnic backgrounds, but it is unclear if the incidence of infection is greater, the prognosis is worse once infected, or both.⁵⁻⁸ Patients from Asian, Black and other Minority Ethnic backgrounds are reported to have accounted for 34% of COVID-19 admissions to intensive care (ICU) in England, Wales and Northern Ireland, compared with only 12% for viral pneumonia (2017-2019), and a higher than expected proportion have required organ support and have died.⁹ Of NHS staff known to have died with COVID-19, 63 (63%) were from an ethnic minority background.¹⁰ In the USA, the mortality rate for Black African Americans is quoted as 2.4 times higher than the rate for White Americans.¹¹ It is unclear the extent to which differences in socioeconomic circumstances or pre-existing conditions may account for some or all such differences in risk.¹²

Ethnicity and socioeconomic position were associated with poor health outcomes during previous pandemics.¹³ The effect of ethnicity on outcome for pandemic influenza 2009 was explained in part by socioeconomic status.¹⁴ Ethnic minority groups may be at greater risk of infection, severe disease, and poor outcomes for multiple reasons. These include socioeconomic conditions that increase risks of transmission and vulnerability, such as overcrowded housing,¹⁵ employment in essential frontline occupations,¹⁶ poverty, healthcare seeking behaviours,¹⁷ and reliance on public transport.¹⁸ People living in more deprived areas in the UK have experienced COVID-19 mortality rates more than double those living in less deprived areas.¹⁹

Theoretical biological mechanisms for differences by ethnicity include susceptibility due to common comorbid illnesses such as diabetes and cardiovascular disease,^{20,21} or genetic factors associated with susceptibility or disease progression. Susceptibility to infectious disease,²² including influenza,²³ and viral pneumonia,²⁴ is strongly heritable, and early reports have suggested this is also true for COVID-19.²⁵ Genetic variations in factors affecting viral entry²⁶ or the host immune response²⁷ may lead to differential susceptibility according to genetic heritage. Behavioural differences by ethnicity may also account for variability in healthcare outcomes.

In the wake of the A/H1N1pdm2009 influenza pandemic, the Clinical Characterisation Protocol UK (CCP-UK) for Severe Emerging Infection was developed by the International Severe Acute Respiratory and emerging Infections Consortium (ISARIC) and the World Health Organisation. As the pandemic potential of SARS-CoV-2 became apparent, the CCP was reactivated in the UK on 17th January 2020, in time to recruit the early patients admitted to hospitals with COVID-19.⁴

Our aim was to assess how prognosis in relation to critical care admission, use of invasive mechanical ventilation (IMV), and in-hospital mortality differs by ethnicity in the CCP-UK cohort of hospitalised patients with COVID-19.

Methods

Study design and setting

ISARIC CCP-UK is an ongoing prospective cohort study involving 260 hospitals in England, Scotland, and Wales (table E1). The protocol, revision history, case report form, information leaflets, consent forms, and detail of the Independent Data and Material Access Committee (IDAMAC) are available at <https://isaric4c.net>. The study was approved by the South Central - Oxford C Research Ethics Committee in England (Ref: 13/SC/0149), and by the Scotland A Research Ethics Committee (Ref: 20/SS/0028). The STROBE guidelines were used when reporting.

Participants

Patients were included between 6th February and 8th May 2020 with follow-up until 22nd May 2020. Inclusion criteria were people of all ages who were admitted to hospital in England, Scotland, and Wales, with proven or a high likelihood of infection with SARS-CoV-2 leading to COVID-19 disease. Reverse transcriptase-PCR was the only mode of testing available during the period of study. The decision to test was at the discretion of the clinician attending the patient, and not defined by protocol. The enrolment criterion “high likelihood of infection” reflected that a preparedness protocol cannot assume a diagnostic test will be available for an emergent pathogen. In this activation, site training emphasised importance of only recruiting proven cases.

Data collection

Data were collected directly and extracted from routine healthcare records by research nurses, administrators and medical students, using case report forms on a REDCap (Research Electronic Data Capture, <https://projectredcap.org>) database. With consent, additional biological samples were collected for research. Data regarding illness progression and severity, including location within the hospital (ward vs. critical care) were collected on days 1 (admission/diagnosis), 3, 6, and 9, and on discharge/death, with further daily data collected on consented patients.

Self-reported ethnicity was transcribed from the healthcare record, where it had been recorded along with other demographics on admission. The care report form used an internationally applicable ethnicity definition with people categorised into East Asian, South Asian, West Asian, Black, White, Latin American, Aboriginal/First Nations and Other ethnic minority.²⁸ For the purposes of this analysis, the collective Ethnic Minorities were collapsed into South Asian, East Asian, Black, Other Ethnic Minority (West Asian, Arab, Latin American, Aboriginal/First Nations, Other) and White, based on frequency (table E2).

Recorded comorbidities were asthma, diabetes (type 1 and type 2), chronic cardiac disease (excluding hypertension), chronic haematologic disease, chronic kidney disease, chronic neurological disorder, chronic pulmonary disease (excluding asthma), dementia, HIV/AIDS, malignancy, malnutrition, mild liver disease, moderate/severe liver disease, clinician-assigned obesity, rheumatologic disorder, and smoking.

As the home address was not available, the Index of Multiple Deprivation (IMD) and equivalents in Scotland and Wales, were used to create a hospital-weighted average deprivation score based on the admission population. Definitions varied between nations to accommodate data differences. For patients admitted in England, a weighted deprivation measure was aggregated by hospital catchment (admissions by Lower Layer Super Output Areas [LSOAs] determined from the Hospital Episode Statistics dataset and weighted by the relative contribution of each LSOA). For Wales, an average Welsh Index of Multiple Deprivation (WIMD) was constructed for each hospital by averaging the 2019 version of the WIMD from all admissions to hospital with a pneumonia or flu diagnosis in 2019 (ICD10: J09-J18). For Scotland, a weighted average of the Scottish Index of Multiple Deprivation (2018 mid-year data-zone estimates) was made at the hospital level and

weighted for all non-elective admissions. To account for differences in the construction of these indices across the devolved nations, measures were centred, standardised, and incorporated with a random gradient into models.

Outcomes

The main outcomes were admission to a critical care facility (Intensive Care Unit [level 3] or High Dependency Unit [level 2]), use of IMV, and in-hospital mortality. To avoid potential bias in the assessment of outcomes where the sickest patients have the longest hospital stays, patients admitted to hospital between 8th May and 22nd May 2020 were excluded from outcome analyses.

As a measure of disease severity at admission, the National Early Warning Score 2 (NEWS2) was used.²⁹ NEWS2 distributions were compared between groups and medians compared using the Mann-Whitney U test. Time from symptom onset to admission was modelled using Cox Proportional Hazards regression.

Statistical analyses

Continuous data are presented as mean (standard deviation), or median (interquartile range) if substantially non-normally distributed. Binary data are presented as frequency (%). Binomial confidence intervals for proportions were calculated using Wilson's method.³⁰ For univariable comparisons, we used Welch's t, ANOVA, Mann-Whitney U, or Kruskal-Wallis tests according to data distribution. Categorical data were compared using chi-squared tests.

Our modelling strategy was informed by a putative causal graphical model (figure E1). Critical care admission and IMV use were modelled using hierarchical logistic regression models, adjusting for patient characteristics at level 1 and hospital of treatment at level 2. Baseline models included an adjustment for age, sex, and location. IMD was aggregated at hospital level, centred, and standardised, and allowed to vary as a random gradient by nation (level 3). Associations between ethnicity and comorbidity were modelled individually using logistic regression.

Survival was modelled using Cox Proportional Hazards regression, with alternative methods used for sensitivity analyses. Symptom onset was considered time zero. In the primary approach, discharge from hospital was considered an absorbing state (once discharged, patients were considered no longer at risk of death). Discharged patients were not censored and included in the risk set until the end of follow-up, thus discharge did not compete with death. Sensitivity analyses employed Fine & Gray's competing risks method.³¹ Hierarchical Cox models included hospital as a random intercept and incorporated deprivation gradients varying by country. Parsimonious criterion-based model building used the following principles: relevant explanatory variables were identified via a causal graph *a priori* for exploration; age, sex, and hospital (random effects) were incorporated in baseline models; interactions were checked at first order level; final model selection was informed by the Akaike Information Criterion (AIC) and c-statistic for logistic regression models and c-statistic for survival models, with appropriate assumptions checked including the distribution of residuals and requirement for proportional hazards. The 30-day model-predicted mortality was calculated from the cumulative hazard for particular subgroups of interest and 95% confidence intervals determined by bootstrap resampling of models (5000 iterations).

A classical approach to mediation analysis in survival models was taken, extending into a three-way decomposition of total effect into direct, pure indirect, and mediated interactive effects.³² *A priori* hypotheses around the contribution of different comorbidities to mortality were explored, with comorbidities associated with ethnicity entered into models sequentially. Associations between exposure and outcome, mediators and exposure, and mediators and outcome were characterised. Logistic regression models were

fitted for mediators and Cox proportional survival models for outcomes, which were then combined. Bootstrapping of estimates was performed to provide 95% confidence intervals (95% CIs), together with a bootstrapped estimate of the mediated proportion of the total effect. All mediation models included age and sex as covariates.

Missing data

There were substantial levels of missing data due to the challenges of real-time data collection during a pandemic. Missing data are reported, and patterns of missing data explored (table E9). Multiple imputation of missing values was performed using chained equations for sensitivity analyses. Ten sets, each with 10 iterations, were imputed using available explanatory variables including the outcome. Missing ethnicity data were not imputed, but ethnicity was used in models for imputation. Graphical checks were performed, and imputed sets combined using Rubin's rules.³³

Data were analysed using R (R Core Team v.3.6.3, Vienna, Austria) with packages including *tidyverse*, *finalfit*, *survival*, *cmprsk*, *coxme* and *regmedint*.

Results

Between 6th February and 8th May 2020, we enrolled 34,986 patients admitted to 260 hospitals with COVID-19 in England, Scotland, and Wales. This is estimated to represent 40.4% of all people admitted over that time period to hospital with COVID-19 in England, Scotland, and Wales.

Presentation at hospital

30,693 (88%) had ethnicity data recorded and were considered in the following groups: South Asians (n=1,388, 5%), East Asians (266, 1%), Black (1,094, 4%), Other Ethnic Minorities (2,398, 8%), and Whites (25,547, 83%) (Figure 1).

Ethnic Minorities were younger (White mean 72 years (y) [standard deviation 16], South Asian 60 y [17], East Asian 60 y [17], Black 61 y [17], Other Ethnic Minority 62 y [18]) and more likely to have diabetes (South Asian n=441 [39.4%], East Asian 71 [29.6%], Black 384 [38.8%], Other Ethnic Minority 636 [30.7%]) compared to the White group (5838 [25.1%]) (Table 1).

In unadjusted analysis, Whites were more likely than Ethnic Minorities to have other comorbidities such as chronic cardiac disease (White n=8228 [32.2%], South Asian 266 [19.2%], East Asian 47 [17.7%], Black 162 [14.8%], Other Ethnic Minority 467 [19.5%]); non-asthmatic chronic pulmonary disease (White 4751 [18.6%], South Asian 86 [6.2%], East Asian 18 [6.8%], Black 73 [6.7%], Other Ethnic Minority 218 [9.1%]); and dementia (White 4106 [16.1%], South Asian 60 [4.3%], East Asian 19 [7.1%], Black 74 [6.8%], Other Ethnic Minority 193 [8.0%]).

With age and sex adjustment, ethnic differences in comorbidity persisted with all Ethnic Minorities more likely to have diabetes (type 1 or type 2) and less likely to have non-asthmatic chronic pulmonary disease and obesity than the Whites (Figure 2). No difference in the likelihood of chronic cardiac disease was seen between South Asians and Whites.

No significant differences were seen in the distribution of severity scores (national early warning score 2 [NEWS2]) at admission between ethnic groups (figure 3A and figure 3C). In patients presenting to hospital with COVID-19, no differences were seen in the time from symptom onset to admission between ethnic groups (figure 3B) when modelled using Cox proportion hazards methods for the relative hazard of admission (figure 3D).

Ethnicity, critical care admission and invasive mechanical ventilation

Overall, 4,353 patients (14%) were admitted to a critical care facility (figure 4A). On unadjusted analysis, White individuals (n=3,258, 13%) were less likely to be admitted to critical care than South Asian (293, 21%), East Asian (64, 24%), Black (241, 22%), or Other Ethnic Minority (497, 21%) people.

In models accounting for age, sex, and location (figure 5B), these associations persisted in South Asian (odds ratio 1.28, 95% confidence interval 1.09 to 1.59), Black (1.36, 1.14 to 1.62), and Other Ethnic Minority (1.29, 1.13 to 1.47) (table E3) groups. No meaningful change in these estimates was seen with the sequential introduction of potential mediators including deprivation and comorbidities (diabetes, obesity, chronic cardiac disease, non-asthmatic chronic pulmonary disease, chronic kidney disease, and dementia) (figure 5C).

Similar findings were found in analyses of invasive mechanical ventilation (Figure E2, table E3).

Ethnicity and survival from COVID-19

In analyses not accounting for age and sex differences between ethnic groups, in-hospital mortality occurred more frequently in the White group compared to the Ethnic Minorities (figure 4C, 6A). A sensitivity analysis excluding patients admitted in the most recent four weeks showed the same results (figure 4D).

However, in models accounting for age, sex, and location, evidence of higher mortality was seen in South Asian (hazard ratio 1.19, 1.05 to 1.36), but not East Asian (1.00, 0.74 to 1.35), Black (1.05, 0.91 to 1.21) or Other Ethnic Minority (0.99, 0.89 to 1.10) groups, compared to White group (figure 6B; table E5). No interaction was seen between ethnicity and age nor ethnicity and sex. Similar results were seen in alternative models of competing risks (table E10) and in models replaced in the ten multiple imputation sets (figure E3).

All associated comorbidities were explored as potential mediators of the apparent association between South Asian ethnicity and an increased hazard of death. Comorbidities associated with death in age- and sex-adjusted analysis were explored (table E11 to E15). Of these, diabetes (hazard ratio 2.29, 1.99 to 2.62, $P < 0.001$) and chronic kidney disease (1.66, 1.38 to 2.00, $P < 0.001$) had strong positive associations with South Asian ethnicity (figure 2).

A significant mediation effect (total natural indirect) of diabetes was found (hazard ratio 1.03, 1.02 to 1.04, $P < 0.001$; table E16) representing 17.8% (8.9 to 65.7, $P = 0.009$; table E17) of the total effect of South Asian ethnicity on mortality (figure 7). Chronic kidney disease did not contribute significantly beyond the addition of diabetes to the model. Finally, to explore the relationship between age, diabetes and South Asian ethnicity, a Cox model was fitted including a three-way interaction (Figure 7B). The excess mortality was seen only in older South Asian patients. The association between diabetes and mortality was strongest in younger patients, compared with older. The diabetes effect was as strong in white patients, but the prevalence of diabetes higher in the South Asian group.

Discussion

In this UK prospective cohort of 34,986 patients in hospital with COVID-19, patients from Ethnic Minority backgrounds were more likely to be admitted to critical care and to undergo IMV than White people, despite there being no difference in disease severity at presentation nor duration of symptoms and being substantially younger. South Asians in hospital with COVID-19 were 20% more likely to die than Whites, and 18% of this increased risk was mediated by higher prevalence of diabetes in the South Asian group. The higher risk of death was apparent in older South Asians, while the effect of diabetes was strongest in the young. This has significant implications given the younger age and higher prevalence of diabetes in South Asians admitted to hospital with COVID-19.

The ISARIC WHO CCP is the largest prospective observational cohort study of in-hospital patients with COVID-19 to date, now accounting for 40% of patients admitted to hospitals in the UK. It is the first study to look at in-hospital trajectories between different ethnic groups with COVID-19. We analysed the data with several separate approaches and our findings remained consistent. Our study has demonstrated the vital importance of forward planning and investment in preparedness studies: we were able to collect granular patient data in near-real time, and undertake analyses during the growth phase of the pandemic in the UK, without the delay associated with routine healthcare data linkages.

As this was a cohort of patients already admitted to hospital, we were unable to assess the risk of hospital admission or out-of-hospital mortality, leading to a potential risk of selection/collider bias when assessing mediators. Our analysis also does not take community factors into account such as population infection rates, exposure by employment, dose of exposure, and potential differential infection rates between different ethnic groups. Higher infection rates may be the result of bigger families, multi-generational households, urban residence, domestic overcrowding, and a greater proportion of individuals in direct public-interacting and healthcare occupations. Our analysis was also limited in the measurement of some key measures of health inequality: we did not have individual measures of socioeconomic status and had limited information on health behaviours such as smoking.

While we recruited 40% of patients admitted to hospital with COVID-19, the results may reflect the differential risk in the whole hospital COVID-19 population, however we have no information on the 60% of patients who were not recruited. Our ethnicity definitions were designed to be broad for ease of data collection, but this reduced the ability to capture nuances in ethnicity. This limited our ability to explore these in greater depth, meaning that we may have missed differences within smaller groups. The Other Ethnic Minority group was likely to be heterogeneous. Obesity was defined by physician assignment and recorded as present or absent, rather than as BMI, which may underrepresent obesity in Asian population where obesity is defined as BMI greater than 27.5.³⁴ There were significant missing data given the nature of collecting prospective data at speed during a pandemic, but analyses of imputed datasets reflect those of the complete case analysis.

Characteristics of ethnic groups going into hospital were different: patients from non-White groups were younger and more likely to have diabetes compared with the White group, who were older and had more cardiac and respiratory disease. Variation in healthcare seeking behaviours and access to healthcare amongst more disadvantaged groups could also contribute to differences in the consequences of COVID-19,¹⁸ however, we found no significant difference between ethnic groups in duration or severity of symptoms at hospital presentation in patients who were admitted.

Although the age-adjusted proportions of patients admitted to critical care was higher in the Ethnic Minority group, a large majority of patients admitted to critical care were White (White 75%, South Asian 6.7%, East Asian 1.5%, Black 5.5%, Other Ethnic Minority 11.4%). This is a higher proportion compared with the latest

ICNARC report, which found a two thirds of patients admitted to critical care were White.¹⁹ This may represent a selection bias in our cohort, or differences in definitions of critical care.

As reported in our earlier analysis, extreme elderly age and the presence of comorbidities were associated with a lower likelihood of critical care admission, suggesting advanced care planning decisions on the ward.⁴ While this may have disproportionately affected patients in the White group who were significantly older with more cardiac and respiratory disease, the increased likelihood of critical care admission and IMV in Ethnic Minorities persisted after adjustment. This may reflect an increased severity of disease in multiple Ethnic Minority populations, but falling short of significantly increased mortality in many.

Our finding of significantly increased mortality in the South Asian group is in agreement with reports which found that people from Asian and Black ethnic groups were at increased risk of in-hospital death from COVID-19, and that this was only partially attributable to pre-existing clinical risk factors or deprivation.^{8,18,35} We add an important and detailed picture of differences in characteristics of ethnic groups in hospital. The prevalence of diabetes is also striking at around 40% in South Asian and Black patients and the mediation of the mortality effect by diabetes is important. This appeared independent of chronic cardiac disease and chronic kidney disease and is likely to be due to the multitude of other negative health associations carried by the condition.

What of the increased mortality in South Asian patients left unexplained? Our models did not account well for wide socioeconomic determinants of outcome, as well as nuance in the comorbidity data. Socioeconomic inequalities contribute to differential exposure to infection, differential vulnerability following infection, and differential consequences of infection as well as control measures.³⁶ In other studies, at Local Authority district level, the districts with a greater proportion of residents from non-White groups experienced higher COVID19 mortality rates, as did districts with a greater proportion of residents experiencing deprivation from low income.¹⁹ We did not find that deprivation at the level of the hospital was a significant predictor of time to hospital admission or severity of illness in hospital, however, aggregation of deprivation at hospital level masks individual socioeconomic effects. Future studies should examine the disaggregated effects where possible.

Why was excess mortality not apparent among our Black population, as others have reported?⁸ It seems unlikely this group are dying disproportionately in community compared with hospital settings, so any true effect should be visible in our cohort. Given that our study includes 40% of the UK hospitalised population, selection bias is clearly a possibility. On the other hand, it is possible that the estimates in our study are correct, given that some population-level reports rely on older census data as the denominator. Our on-going work includes linking our dataset widely to better understand the place of the in-hospital stay as part of the full patient journey.

Finally, a subset of our cohort has consented to biological studies that can investigate potential bio-mechanistic differences in outcomes by ethnicity. Susceptibility to influenza²³ is heritable, and it seems likely that this is also true for COVID-19.²⁵ Ethnicity correlates with particular genetic polymorphisms which may link to important biological mechanisms mediating disease susceptibility and severity, for instance, relating to ACE2 receptor expression and SARS-CoV-2 cell entry.²⁶

Our findings have important implications for policy makers and researchers. Given we illustrate increased disease severity across multiple hospitalised Ethnic Minority populations, ethnicity may need to be considered in the prioritisation of treatment and future vaccination policy. South Asians are over-represented in frontline key worker and other direct public-interacting occupations,³⁷ and policies should consider this factor in refining guidance and advice for at-risk individuals, accounting for age and comorbidity. On-going work to devise risk assessment tools to help mitigate the likelihood of SARS-CoV-2

infection and manage the easing of individual lockdown restrictions needs to consider the effects of ethnicity on outcomes. Careful consideration also needs to be given in constructing these tools to the weighting of mediators of risk such as diabetes. Further research integrating primary and secondary care data is key to improving the understanding of other drivers of poor outcomes in Ethnic Minority patients.

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Table 1. Characteristics of patients admitted to hospital with COVID-19 during between 6th February 2020 and 8th May 2020 by ethnic group. P-value is for a comparison across ethnic groups excluding missing data and is a chi-squared test for categorical variables and an F-test for continuous variables. Data n (%) unless otherwise stated. IMD, index of multiple deprivation. 4,293 missing ethnicity (see table E9).

Ethnicity		White	South Asian	East Asian	Black	Other Ethnic Minority	p
Total N (%)		25547 (83.2)	1388 (4.5)	266 (0.9)	1094 (3.6)	2398 (7.8)	
Age on admission (years)	Mean (SD)	72.3 (15.6)	59.8 (17.4)	60.3 (17.3)	61.4 (17.3)	62.4 (17.8)	<0.001
	18-39	1025 (4.0)	180 (13.0)	32 (12.0)	126 (11.5)	275 (11.5)	
	40-49	1308 (5.1)	243 (17.5)	43 (16.2)	140 (12.8)	336 (14.0)	
	50-59	2962 (11.6)	248 (17.9)	57 (21.4)	273 (25.0)	432 (18.0)	
	60-69	3928 (15.4)	286 (20.6)	43 (16.2)	197 (18.0)	431 (18.0)	
	70-79	6319 (24.7)	222 (16.0)	46 (17.3)	143 (13.1)	430 (17.9)	
	80+	10005 (39.2)	209 (15.1)	45 (16.9)	215 (19.7)	494 (20.6)	
	Sex at Birth	Male	14684 (57.5)	887 (63.9)	174 (65.4)	602 (55.0)	1426 (59.5)
Female		10778 (42.2)	500 (36.0)	91 (34.2)	490 (44.8)	964 (40.2)	
(Missing)		85 (0.3)	1 (0.1)	1 (0.4)	2 (0.2)	8 (0.3)	
Diabetes	No	17447 (74.9)	679 (60.6)	169 (70.4)	606 (61.2)	1435 (69.3)	<0.001
	Yes	5838 (25.1)	441 (39.4)	71 (29.6)	384 (38.8)	636 (30.7)	
Obesity	No	18952 (74.2)	910 (65.6)	207 (77.8)	804 (73.5)	1718 (71.6)	0.201
	Yes	2382 (9.3)	114 (8.2)	19 (7.1)	123 (11.2)	210 (8.8)	
	(Missing)	4213 (16.5)	364 (26.2)	40 (15.0)	167 (15.3)	470 (19.6)	
Chronic cardiac disease	No	15403 (60.3)	867 (62.5)	195 (73.3)	841 (76.9)	1629 (67.9)	<0.001
	Yes	8228 (32.2)	266 (19.2)	47 (17.7)	162 (14.8)	467 (19.5)	
	(Missing)	1916 (7.5)	255 (18.4)	24 (9.0)	91 (8.3)	302 (12.6)	
Chronic pulmonary disease	No	18821 (73.7)	1043 (75.1)	223 (83.8)	929 (84.9)	1857 (77.4)	<0.001
	Yes	4751 (18.6)	86 (6.2)	18 (6.8)	73 (6.7)	218 (9.1)	
	(Missing)	1975 (7.7)	259 (18.7)	25 (9.4)	92 (8.4)	323 (13.5)	
Asthma	No	20079 (78.6)	945 (68.1)	207 (77.8)	870 (79.5)	1803 (75.2)	0.108
	Yes	3358 (13.1)	188 (13.5)	34 (12.8)	129 (11.8)	281 (11.7)	
	(Missing)	2110 (8.3)	255 (18.4)	25 (9.4)	95 (8.7)	314 (13.1)	
Chronic kidney disease	No	19201 (75.2)	931 (67.1)	218 (82.0)	831 (76.0)	1796 (74.9)	<0.001
	Yes	4208 (16.5)	193 (13.9)	26 (9.8)	170 (15.5)	286 (11.9)	
	(Missing)	2138 (8.4)	264 (19.0)	22 (8.3)	93 (8.5)	316 (13.2)	
Moderate/severe liver disease	No	22739 (89.0)	1102 (79.4)	237 (89.1)	976 (89.2)	2012 (83.9)	0.221
	Yes	460 (1.8)	13 (0.9)	2 (0.8)	17 (1.6)	37 (1.5)	
	(Missing)	2348 (9.2)	273 (19.7)	27 (10.2)	101 (9.2)	349 (14.6)	
Mild liver disease	No	22798 (89.2)	1096 (79.0)	233 (87.6)	979 (89.5)	2006 (83.7)	0.719
	Yes	344 (1.3)	18 (1.3)	6 (2.3)	13 (1.2)	32 (1.3)	
	(Missing)	2405 (9.4)	274 (19.7)	27 (10.2)	102 (9.3)	360 (15.0)	
Chronic neurological disorder	No	20243 (79.2)	1045 (75.3)	221 (83.1)	928 (84.8)	1870 (78.0)	<0.001
	Yes	3032 (11.9)	78 (5.6)	21 (7.9)	68 (6.2)	183 (7.6)	
	(Missing)	2272 (8.9)	265 (19.1)	24 (9.0)	98 (9.0)	345 (14.4)	
Malignancy	No	20665 (80.9)	1072 (77.2)	229 (86.1)	915 (83.6)	1910 (79.6)	<0.001
	Yes	2530 (9.9)	48 (3.5)	9 (3.4)	85 (7.8)	144 (6.0)	
	(Missing)	2352 (9.2)	268 (19.3)	28 (10.5)	94 (8.6)	344 (14.3)	
Chronic haematologic disease	No	22195 (86.9)	1085 (78.2)	232 (87.2)	922 (84.3)	1971 (82.2)	<0.001

	Yes	975 (3.8)	33 (2.4)	8 (3.0)	76 (6.9)	80 (3.3)	
	(Missing)	2377 (9.3)	270 (19.5)	26 (9.8)	96 (8.8)	347 (14.5)	
AIDS/HIV	No	22931 (89.8)	1092 (78.7)	239 (89.8)	947 (86.6)	2020 (84.2)	<0.001
	Yes	74 (0.3)	4 (0.3)	1 (0.4)	34 (3.1)	9 (0.4)	
	(Missing)	2542 (10.0)	292 (21.0)	26 (9.8)	113 (10.3)	369 (15.4)	
Rheumatologic disorder	No	20412 (79.9)	1042 (75.1)	230 (86.5)	936 (85.6)	1899 (79.2)	<0.001
	Yes	2717 (10.6)	67 (4.8)	8 (3.0)	62 (5.7)	141 (5.9)	
	(Missing)	2418 (9.5)	279 (20.1)	28 (10.5)	96 (8.8)	358 (14.9)	
Dementia	No	19252 (75.4)	1055 (76.0)	220 (82.7)	933 (85.3)	1872 (78.1)	<0.001
	Yes	4106 (16.1)	60 (4.3)	19 (7.1)	74 (6.8)	193 (8.0)	
	(Missing)	2189 (8.6)	273 (19.7)	27 (10.2)	87 (8.0)	333 (13.9)	
Malnutrition	No	21456 (84.0)	1050 (75.6)	225 (84.6)	948 (86.7)	1922 (80.2)	0.031
	Yes	607 (2.4)	20 (1.4)	5 (1.9)	13 (1.2)	46 (1.9)	
	(Missing)	3484 (13.6)	318 (22.9)	36 (13.5)	133 (12.2)	430 (17.9)	
Smoking	Yes	1266 (5.0)	34 (2.4)	9 (3.4)	27 (2.5)	95 (4.0)	<0.001
	Never Smoked	9271 (36.3)	744 (53.6)	142 (53.4)	609 (55.7)	1133 (47.2)	
	Former Smoker	6310 (24.7)	126 (9.1)	34 (12.8)	114 (10.4)	332 (13.8)	
	(Missing)	8700 (34.1)	484 (34.9)	81 (30.5)	344 (31.4)	838 (34.9)	
Prior immunosuppression	No	20734 (81.2)	1027 (74.0)	216 (81.2)	907 (82.9)	1906 (79.5)	0.003
	Yes	2329 (9.1)	88 (6.3)	15 (5.6)	86 (7.9)	173 (7.2)	
	(Missing)	2484 (9.7)	273 (19.7)	35 (13.2)	101 (9.2)	319 (13.3)	
Prior infection treatment	No	18342 (71.8)	892 (64.3)	187 (70.3)	817 (74.7)	1672 (69.7)	0.193
	Yes	4659 (18.2)	216 (15.6)	44 (16.5)	176 (16.1)	389 (16.2)	
	(Missing)	2546 (10.0)	280 (20.2)	35 (13.2)	101 (9.2)	337 (14.1)	
Deprivation (IMD)	Mean (SD)	0.1 (0.8)	0.1 (0.8)	0.2 (0.8)	0.1 (0.7)	0.3 (0.7)	<0.001

Figure 1. Patient inclusion stratified by ethnicity. Critical care use, invasive mechanical ventilation, and mortality in all included patients.

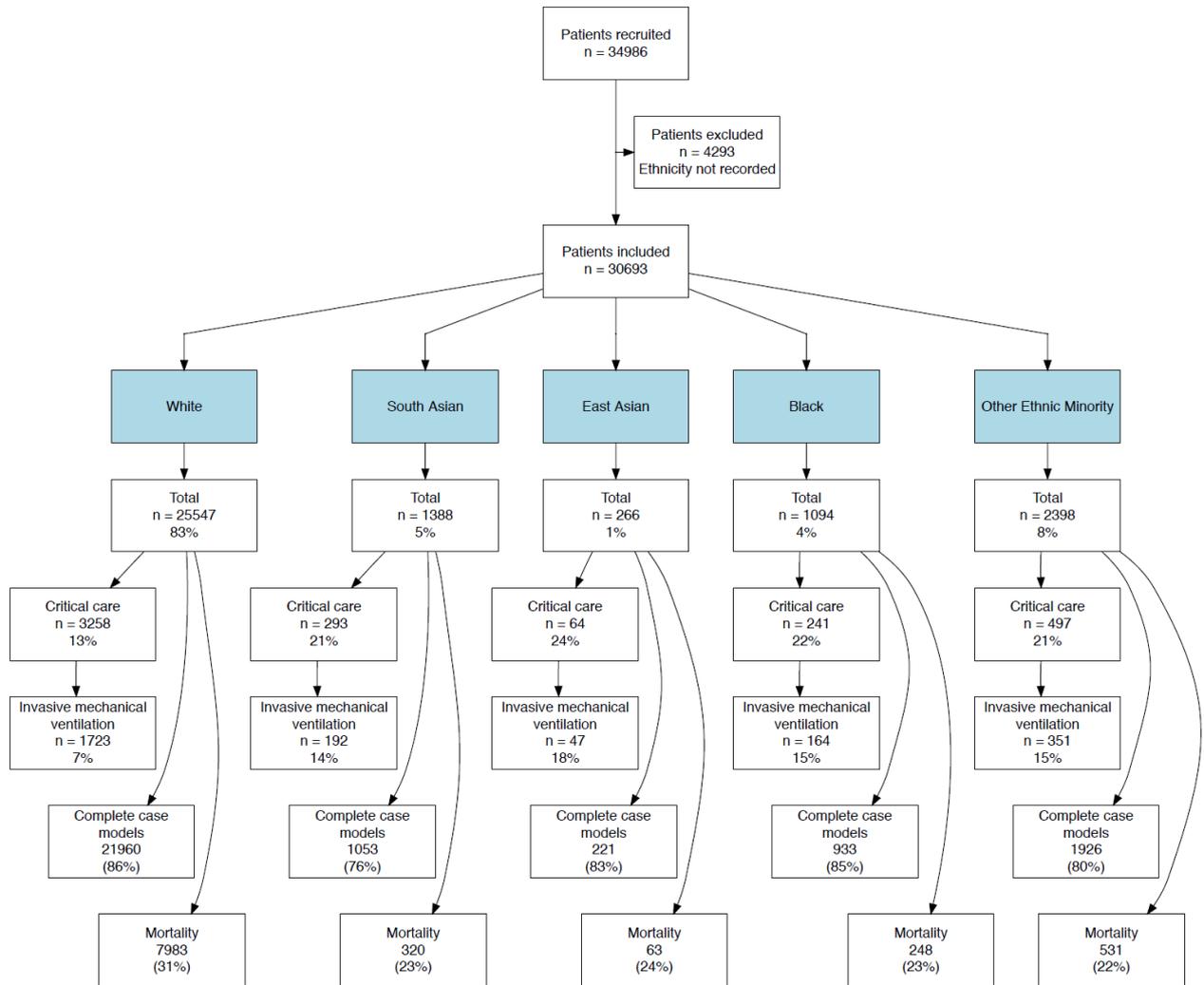


Figure 2. Age and sex adjusted association between ethnicity and comorbidity for patients in hospital with COVID-19. Hierarchical logistic regression models of complete data. Data are odds ratios (95% confidence interval, P-value).

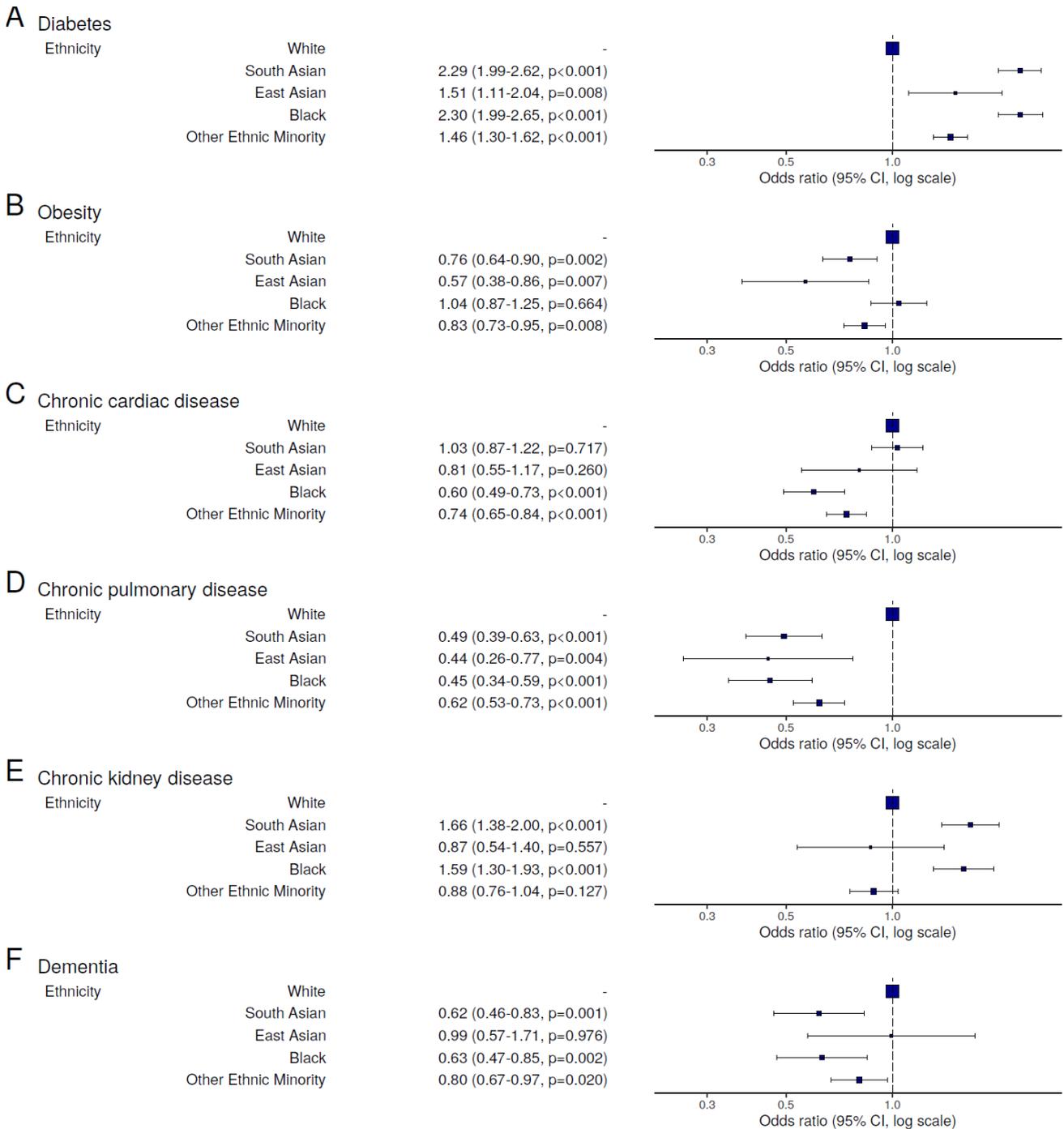


Figure 3. Severity score at admission and time from symptom onset to admission by ethnicity. National early warning score 2 (NEWS2) distribution (A) and proportion with NEWS2 >2 (C) by ethnicity. In patients presenting to hospital with COVID-19, distribution of time from symptom onset to admission (B) and Cox proportion hazards model of time to admission showing relative hazard for admission by ethnicity adjusted for age and sex (D).

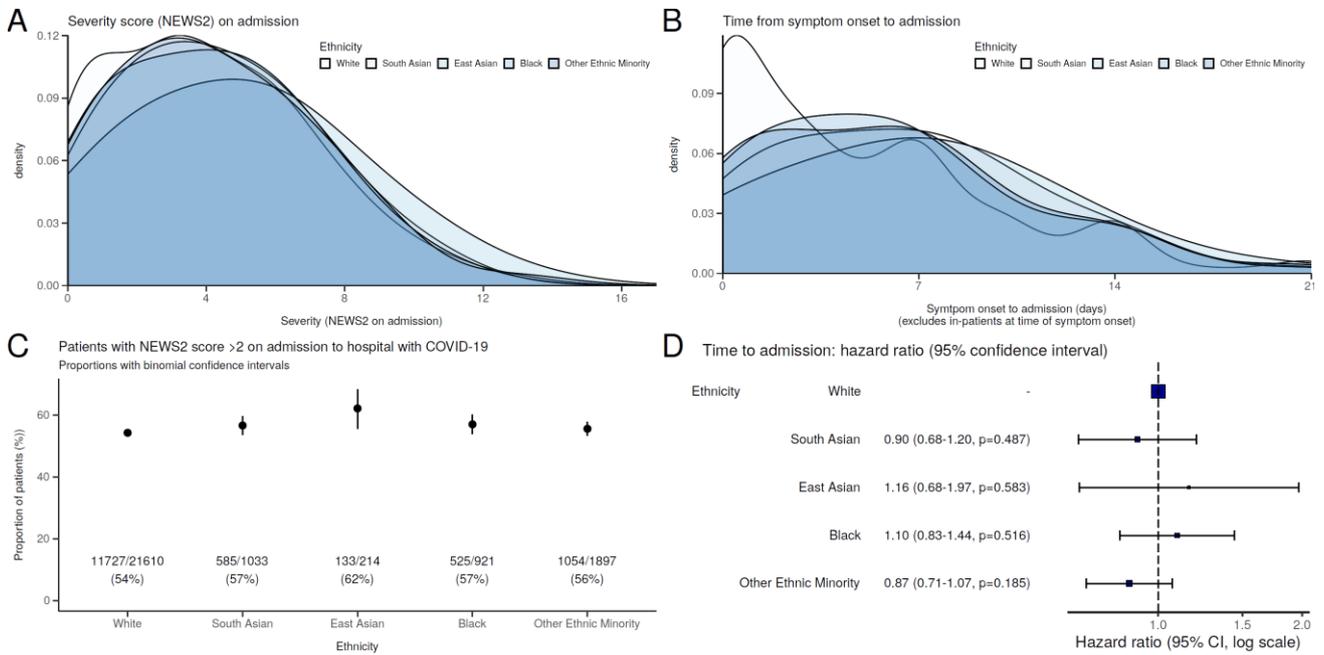


Figure 4. Unadjusted critical care admission (A), use of invasive mechanical ventilation (B), case fatality in patients with at least 2 weeks follow-up (C) and with at least 4 weeks follow-up (D) by ethnicity. Proportions from crude counts with no adjustment for age, sex, or clustering. Binomial confidence intervals. Numbers are patient counts per group / total cohort (%). IMV, invasive mechanical ventilation.

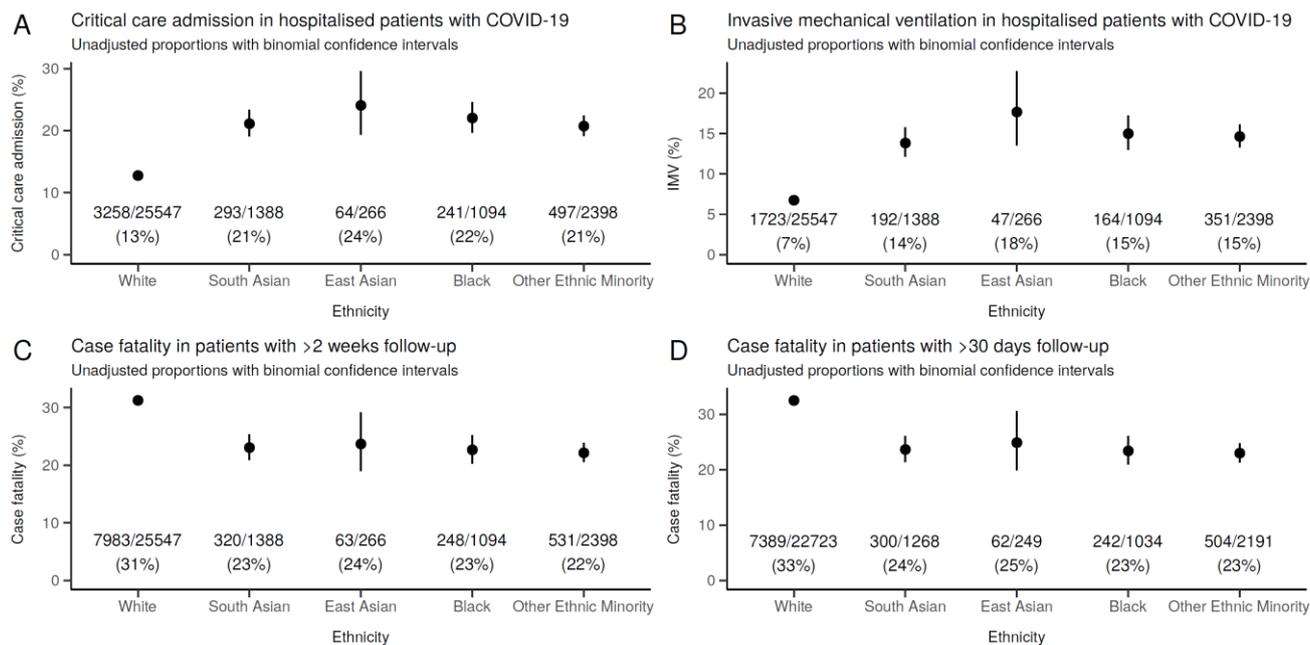
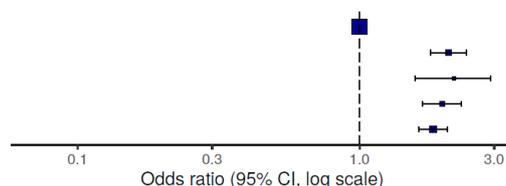


Figure 5. Critical care admission by ethnic group. (A) Univariable logistic regression model of ethnicity alone. (B) Hierarchical logistic regression model adjusting for age, sex and accounting for clustering in hospitals. (C) Hierarchical logistic regression models with adjustment for factors potentially mediating critical care admission were explored sequentially. Full model presented here accepting potential bias in comorbidity estimates. Complete case data. See table E3 for full models. Size of point estimate box proportional to population size. IMD, index of multiple deprivation (centred and standardised).

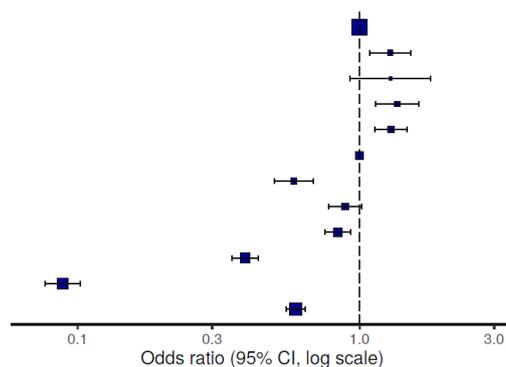
A Critical care admission: ethnicity alone

Ethnicity	White	-
	South Asian	2.07 (1.78-2.39, p<0.001)
	East Asian	2.16 (1.57-2.91, p<0.001)
	Black	1.96 (1.67-2.29, p<0.001)
	Other Ethnic Minority	1.82 (1.62-2.04, p<0.001)



B Critical care: hierarchical baseline

Ethnicity	White	-
	South Asian	1.28 (1.09-1.52, p=0.003)
	East Asian	1.29 (0.92-1.79, p=0.136)
	Black	1.36 (1.14-1.62, p=0.001)
	Other Ethnic Minority	1.29 (1.13-1.47, p<0.001)
Age on admission (years)	50-59	-
	18-39	0.58 (0.50-0.68, p<0.001)
	40-49	0.89 (0.78-1.01, p=0.080)
	60-69	0.84 (0.75-0.93, p=0.001)
	70-79	0.39 (0.35-0.44, p<0.001)
	80+	0.09 (0.08-0.10, p<0.001)
Sex at Birth	Female	0.59 (0.55-0.64, p<0.001)



C Critical care: hierarchical with potential mediators

Ethnicity	White	-
	South Asian	1.32 (1.12-1.56, p=0.001)
	East Asian	1.30 (0.93-1.81, p=0.127)
	Black	1.34 (1.12-1.61, p=0.001)
	Other Ethnic Minority	1.26 (1.10-1.44, p=0.001)
Age on admission (years)	50-59	-
	18-39	0.55 (0.47-0.64, p<0.001)
	40-49	0.86 (0.75-0.98, p=0.026)
	60-69	0.95 (0.86-1.06, p=0.378)
	70-79	0.56 (0.50-0.62, p<0.001)
	80+	0.16 (0.14-0.19, p<0.001)
Sex at Birth	Female	0.56 (0.52-0.61, p<0.001)
Deprivation (IMD)	-	1.03 (0.76-1.39, p=0.873)
Diabetes	Yes	1.00 (0.92-1.09, p=0.969)
Obesity	No	-
	Yes	1.79 (1.61-2.00, p<0.001)
	(Missing)	1.07 (0.92-1.24, p=0.373)
Chronic cardiac disease	Yes	0.64 (0.57-0.71, p<0.001)
Chronic pulmonary disease	Yes	0.71 (0.63-0.80, p<0.001)
Chronic kidney disease	Yes	0.66 (0.57-0.75, p<0.001)
Dementia	Yes	0.26 (0.21-0.33, p<0.001)

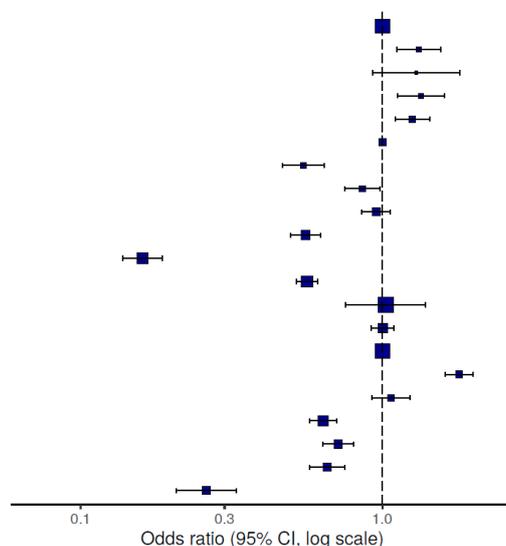


Figure 6. In-hospital survival by ethnic group. Estimates are hazard ratio (95% confidence interval, P-value). (A) Univariable Cox proportional hazards regression model of ethnicity alone. (B) Hierarchical Cox proportional hazards regression model adjusting for age, sex and deprivation, accounting for clustering within hospital. (C) Hierarchical Cox proportional hazards regression models with adjustment for factors potentially mediating survival by ethnic group were explored sequentially. accounting for clustering. Full model presented here accepting potential bias in comorbidity estimates. Complete case data. See table E5 for full models. Size of point estimate box proportional to population size. CI, confidence interval; IMD, index of multiple deprivation (centred and standardised).

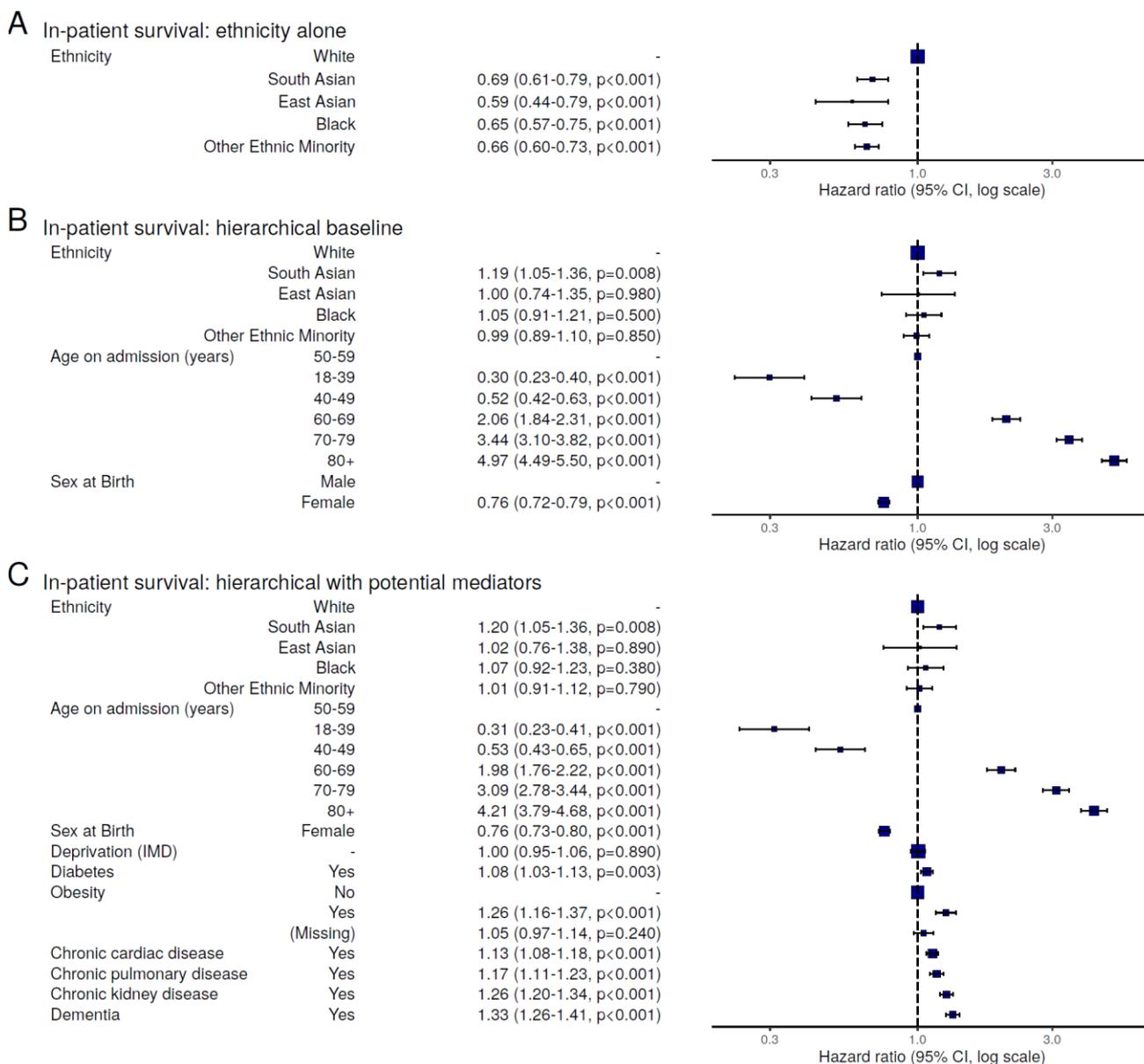
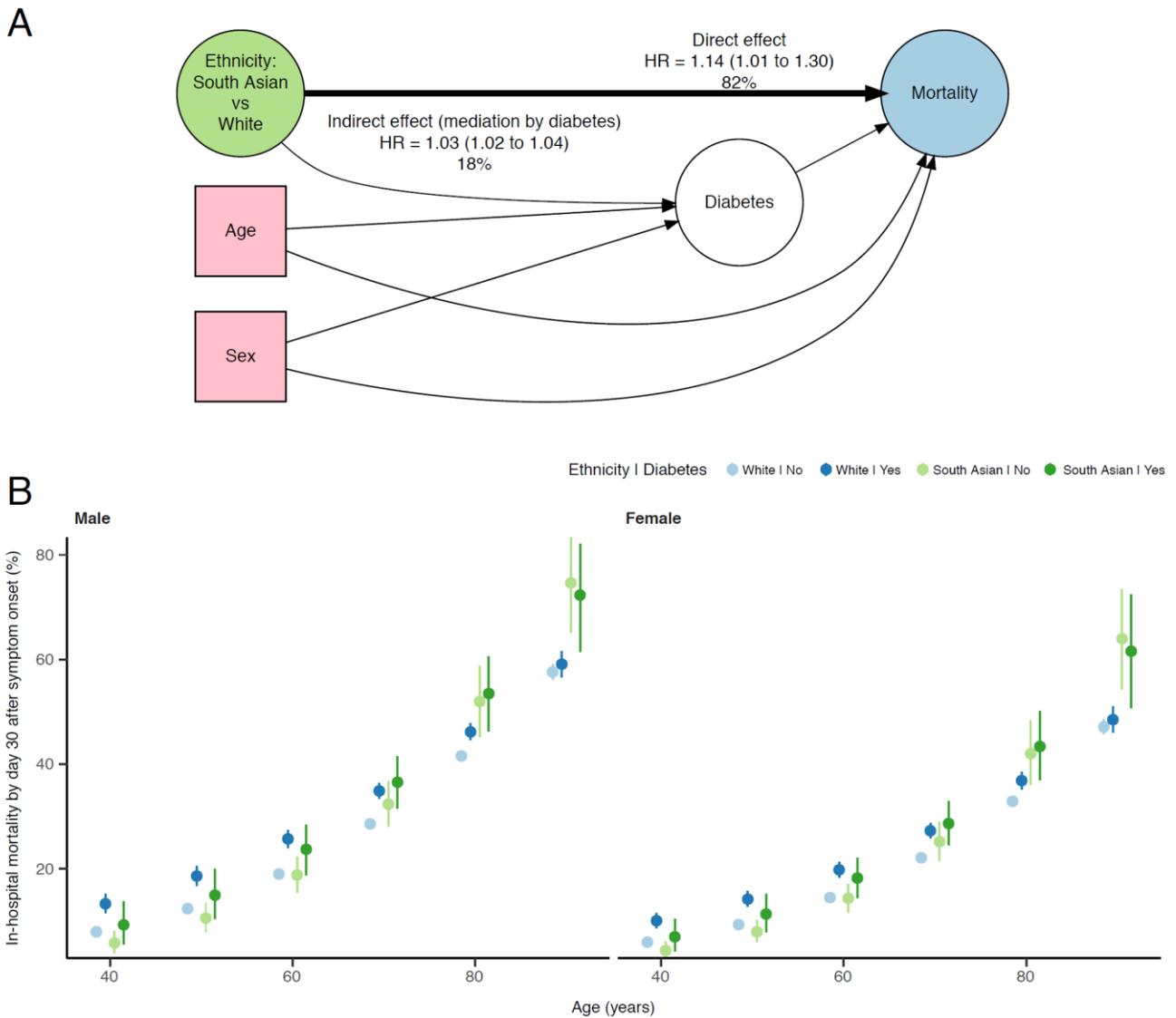


Figure 7. Mediation of apparent excess hazard of death in South Asian group by diabetes. (A) Three-way decomposition mediation analysis was performed for comorbidities associated with ethnicity (Figure 2; table E11 to E15) and outcome (Figure 7). All models included age and sex as covariates. The proportion of the total effect mediated by diabetes for mortality in South Asian patients with bootstrapped 95% CI intervals was determined. (B) Cox proportional hazards model of in-hospital mortality including age, sex, ethnicity, and diabetes status, including interactions between age, ethnicity, and diabetes. The cumulative probability of in-hospital death by day 30 after symptom onset was calculated for the subgroups shown and 95% confidence intervals determined by bootstrap resampling of models. See table E18 for full model.



End matter

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Ethical approval

Ethical approval was given by the South Central – Oxford C Research Ethics Committee in England (Ref: 13/SC/0149), and by the Scotland A Research Ethics Committee (Ref: 20/SS/0028). The study was registered at <https://www.isrctn.com/ISRCTN66726260>.

Contributions

EM Harrison is guarantor and corresponding author for this work, and attests that all listed authors meet authorship criteria and that no others meeting the criteria have been omitted.

Conceptualisation: JK Baillie, J Dunning, G Carson, L Merson, JS Nguyen-Van-Tam, PJM Openshaw, MG Semple. Formal analysis: AB Docherty, C Gamble, EM Harrison, P Horby, L Norman, PJM Openshaw, R Pius, JM Read, MG Semple. Writing original draft: AB Docherty, PJM Openshaw, MG Semple. Writing reviewing and editing JK Baillie, AB Docherty, J Dunning, C Gamble, CA Green, EM Harrison, P Horby, JS Nguyen-Van-Tam, PJM Openshaw, MG Semple, L Sigfrid. Project administration: S Halpin, HE Hardwick, C Gamble, A Ho, KA Holden, J Lee, L Merson, D Plotkin, CD Russell. Investigation: EM Harrison, P Horby, C Gamble, CA Green, A Ho, MG Semple. Supervision: JK Baillie, HE Hardwick, EM Harrison, C Gamble, A Ho, P Horby, PJM Openshaw, MG Semple. Data curation: L Merson, S Halpin, C Jackson. Validation: KA Holden, S Halpin, C Jackson, Funding acquisition: JK Baillie, G Carson, P Horby, PJM Openshaw, MG Semple.

Competing Interests

All authors have completed the ICMJE uniform disclosure form at www.icmje.org/coi_disclosure.pdf and declare:

AB Docherty reports grants from Department of Health and Social Care, during the conduct of the study; grants from Wellcome Trust, outside the submitted work; CA Green reports grants from DHSC National Institute of Health Research UK, during the conduct of the study; PW Horby reports grants from Wellcome Trust / Department for International Development / Bill and Melinda Gates Foundation, grants from NIHR, during the conduct of the study; JS Nguyen-Van-Tam reports grants from Department of Health and Social Care, England, during the conduct of the study; and is seconded to the Department of Health and Social Care, England (DHSC); PJM Openshaw reports personal fees from consultancies and from European Respiratory Society; grants from MRC, MRC Global Challenge Research Fund, EU, NIHR Biomedical Research Centre, MRC/GSK, Wellcome Trust, NIHR (HPRU in Respiratory Infection), and NIHR Senior Investigator

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Data sharing

We welcome applications for data and material access via our Independent Data And Material Access Committee (<https://isaric4c.net>).

Dissemination to participants and related patient and public communities

ISARIC4C has a public facing website and twitter account @CCPUKstudy. We are engaging with print and internet press, television, radio, news, and documentary programme makers. We will explore distribution of findings with The Asthma UK and British Lung Foundation Partnership and take advice from NIHR Involve and GenerationR Alliance Young People's Advisory Groups.

ISARIC Coronavirus Clinical Characterisation Consortium (ISARIC4C) Investigators

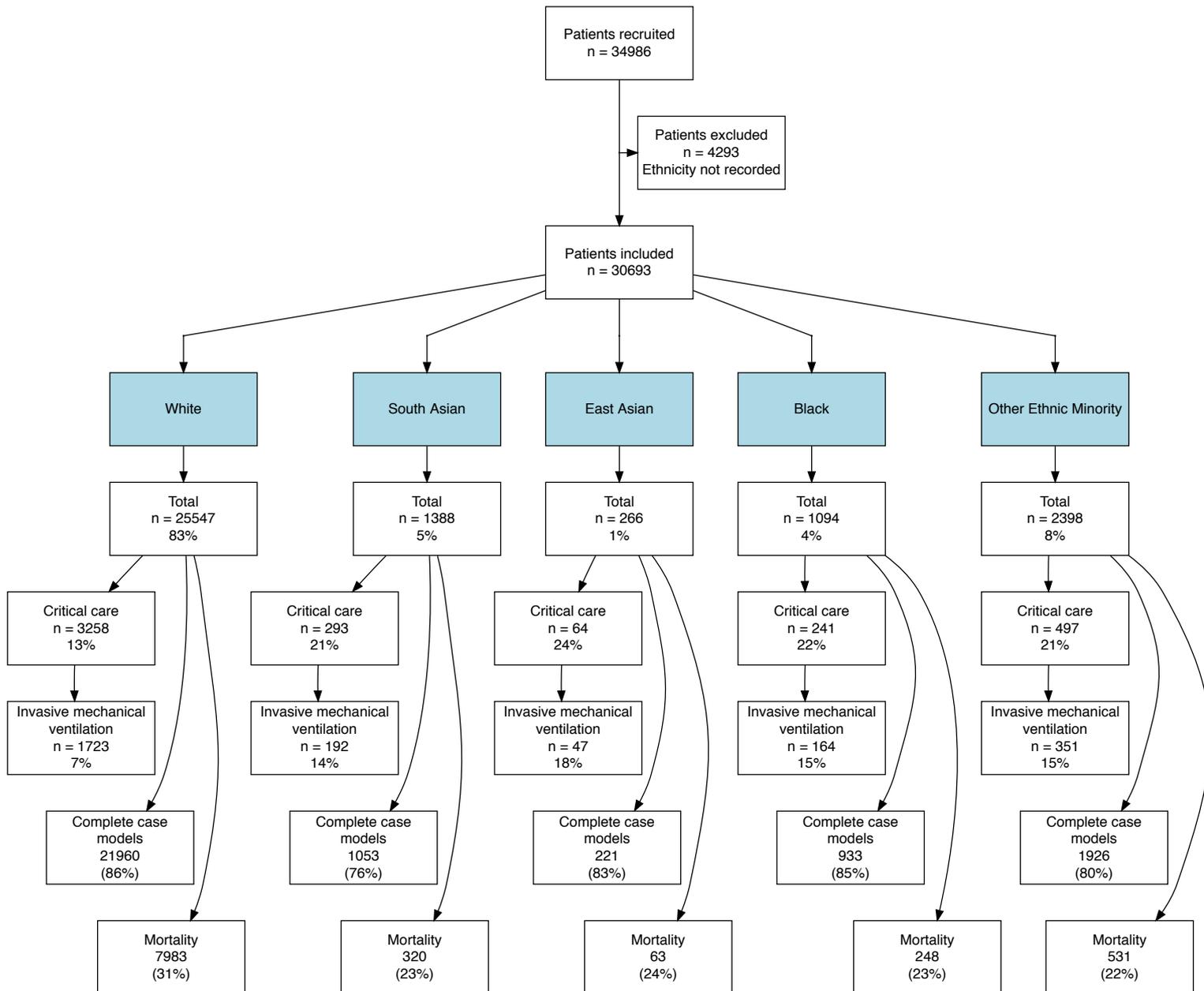
Consortium Lead Investigator J Kenneth Baillie, Chief Investigator Malcolm G Semple. Co-Lead Investigator Peter JM Openshaw. ISARIC Clinical Coordinator Gail Carson. Co-Investigators: Beatrice Alex, Benjamin Bach, Wendy S Barclay, Debby Bogaert, Meera Chand, Graham S Cooke, Annemarie B Docherty, Jake Dunning, Ana da Silva Filipe, Tom Fletcher, Christopher A Green, Ewen M Harrison, Julian A Hiscox, Antonia Ying Wai Ho, Peter W Horby, Samreen Ijaz, Saye Khoo, Paul Klenerman, Andrew Law, Wei Shen Lim, Alexander, J Mentzer, Laura Merson, Alison M Meynert, Mahdad Noursadeghi, Shona C Moore, Massimo Palmarini, William A Paxton, Georgios Pollakis, Nicholas Price, Andrew Rambaut, David L Robertson, Clark D Russell, Vanessa Sancho-Shimizu, Janet T Scott, Tom Solomon, Shiranee Srisikandan, David Stuart, Charlotte Summers, Richard S Tedder, Emma C Thomson, Ryan S Thwaites, Lance CW Turtle, Maria Zambon. Project Managers Hayley Hardwick, Chloe Donohue, Jane Ewins, Wilna Oosthuyzen, Fiona Griffiths. Data Analysts: Lisa Norman, Riinu Pius, Tom M Drake, Cameron J Fairfield, Stephen Knight, Kenneth A Mclean, Derek Murphy, Catherine A Shaw. Data and Information System Manager: Jo Dalton, Michelle Girvan, Egle Saviciute, Stephanie Roberts Janet Harrison, Laura Marsh, Marie Connor. Data integration and presentation: Gary Leeming, Andrew Law, Ross Hendry. Material Management: William Greenhalf, Victoria Shaw, Sarah McDonald. Local Principal Investigators: Kayode Adeniji, Daniel Agranoff, Ken Agwuh, Dhiraj Ail, Ana Alegria, Brian Angus, Abdul Ashish, Dougal Atkinson, Shahedal Bari, Gavin Barlow, Stella Barnass, Nicholas Barrett, Christopher Bassford, David Baxter, Michael Beadsworth, Jolanta Bernatoniene, John Berridge, Nicola Best, Pieter Bothma, David Brealey, Robin Brittain-Long, Naomi Bulteel, Tom Burden, Andrew Burtenshaw, Vikki Caruth, David Chadwick, Duncan Chambler, Nigel Chee, Jenny Child, Srikanth Chukkambotla, Tom Clark, Paul Collini, Graham Cooke, Catherine Cosgrove, Jason Cupitt, Maria-Teresa Cutino-Moguel, Paul Dark, Chris Dawson, Samir Dervisevic, Phil Donnison, Sam Douthwaite, Ingrid DuRand, Ahilanadan Dushianthan, Tristan Dyer, Cariad Evans, Chi Eziefula, Chrisopher Fegan, Adam Finn, Duncan Fullerton, Sanjeev Garg, Sanjeev Garg, Atul

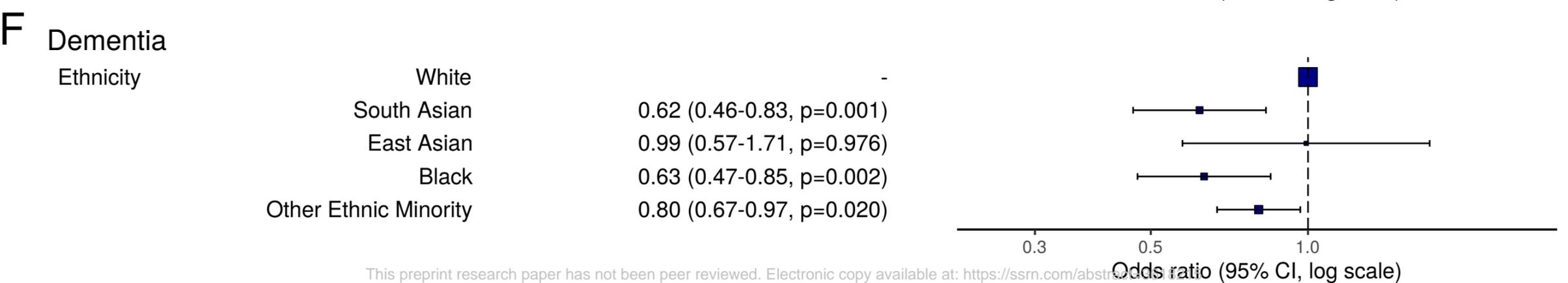
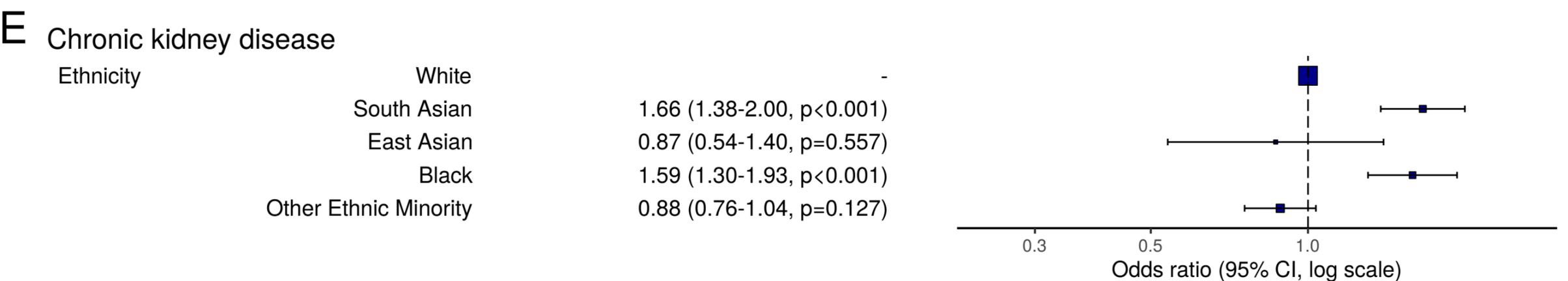
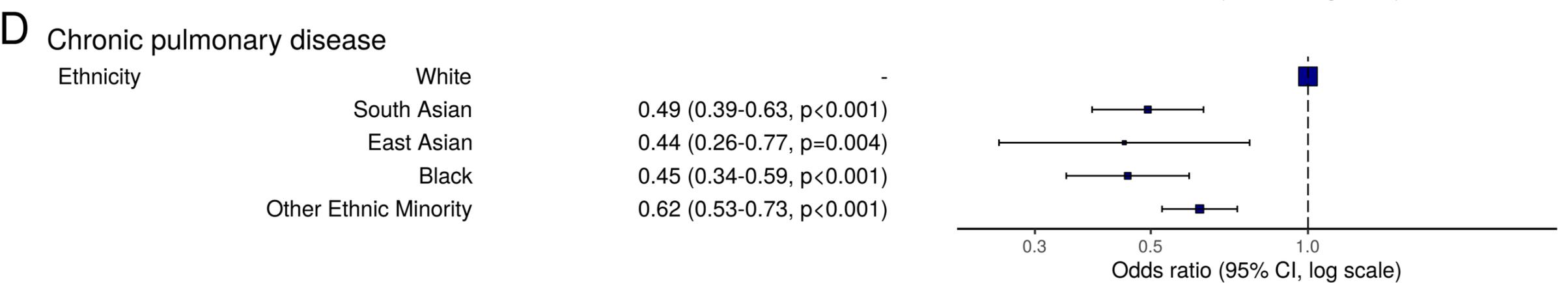
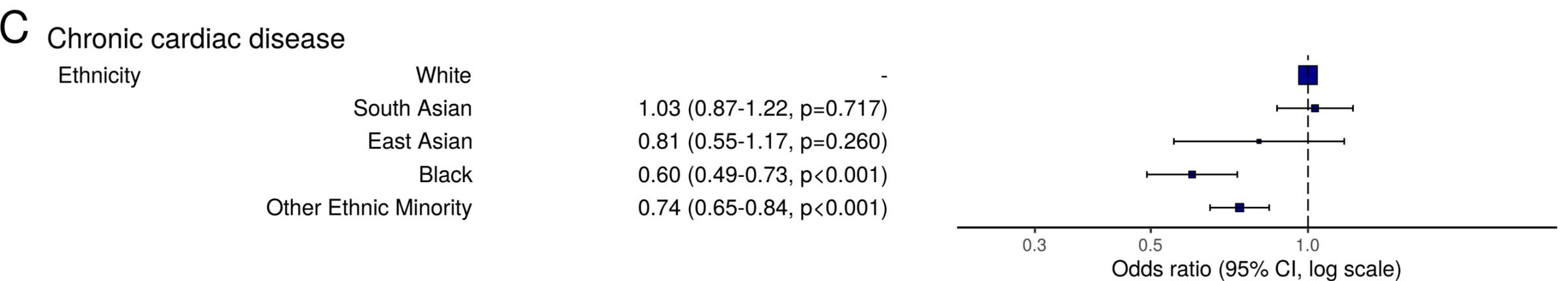
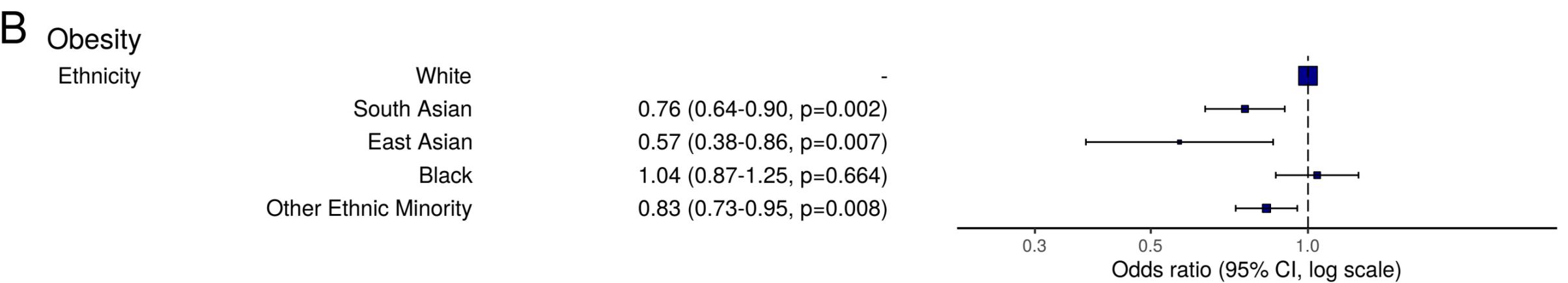
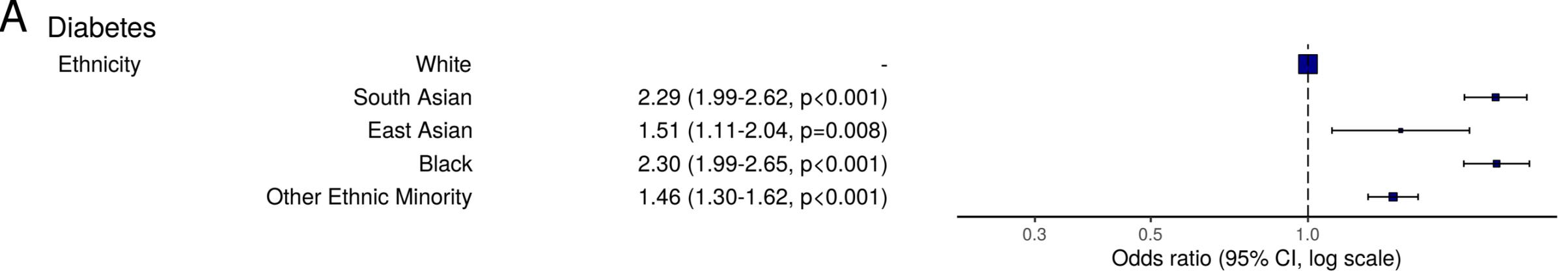
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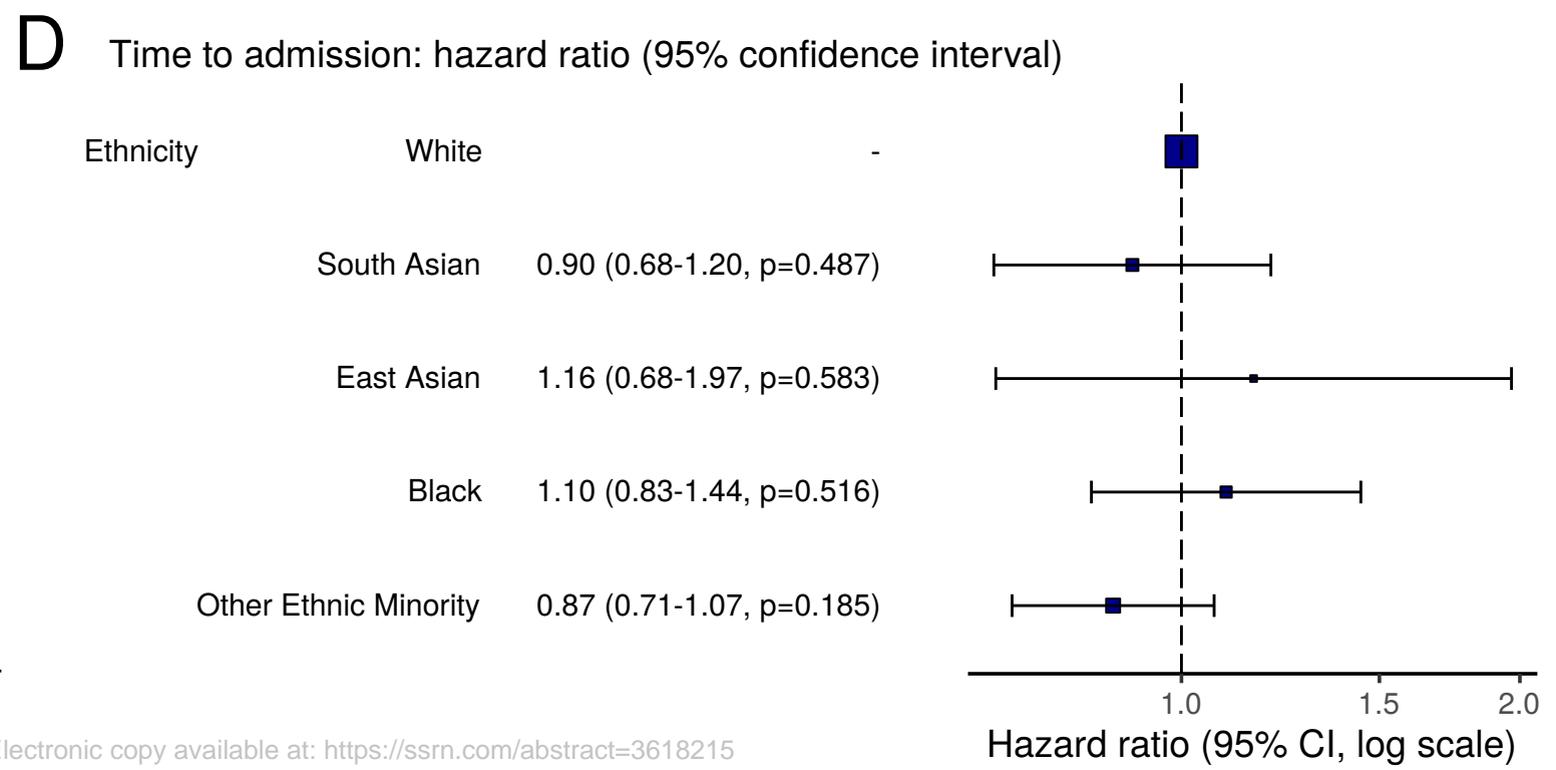
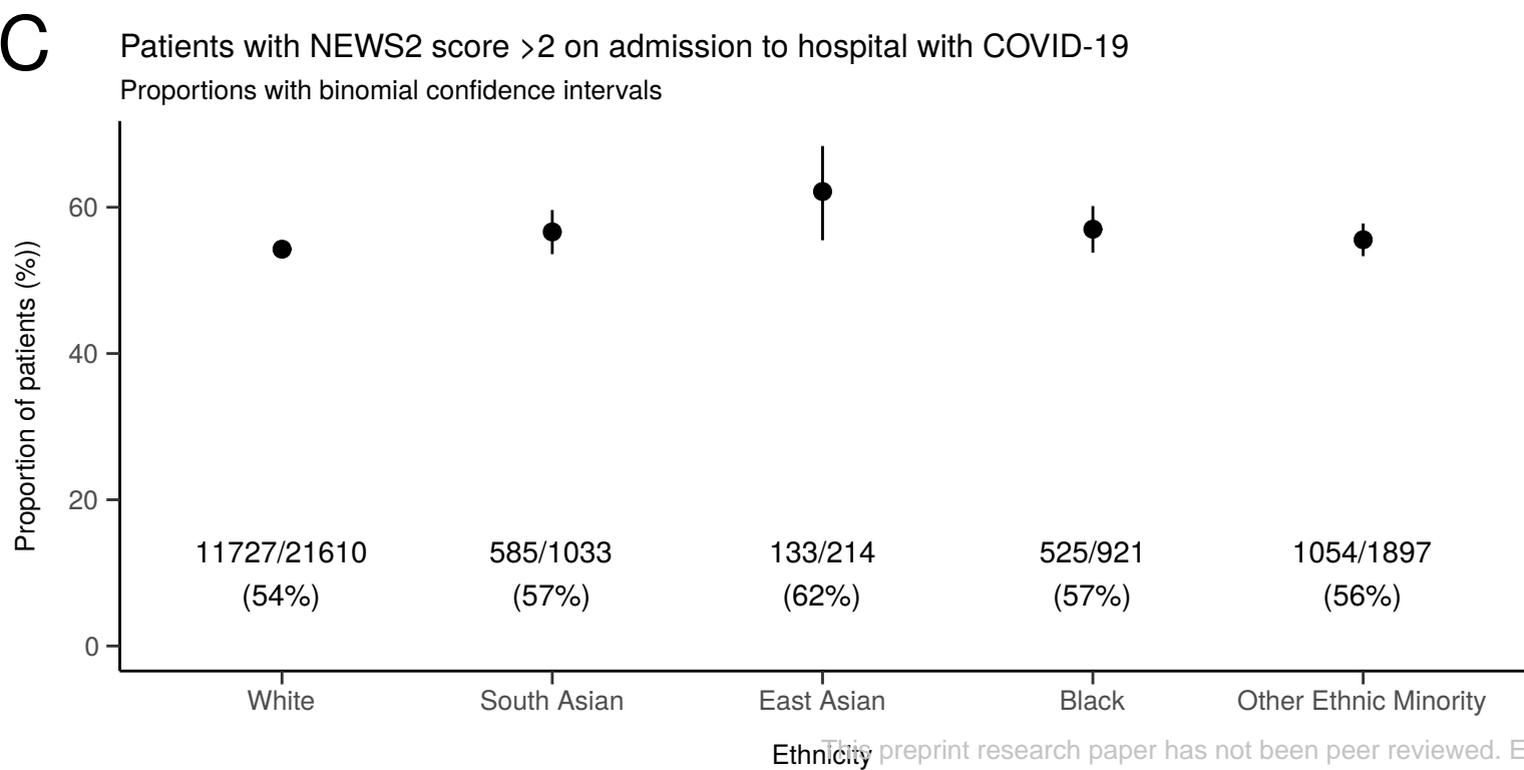
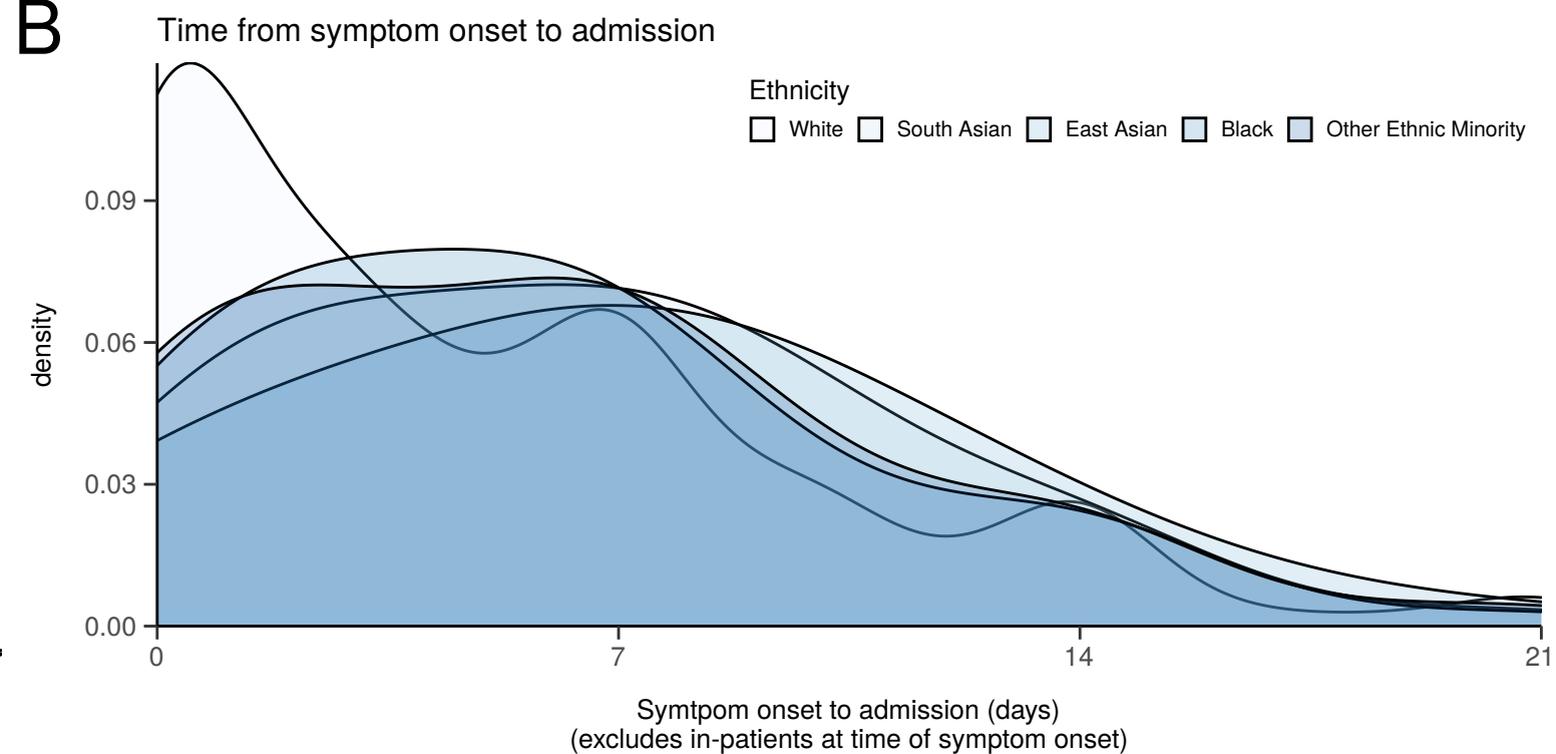
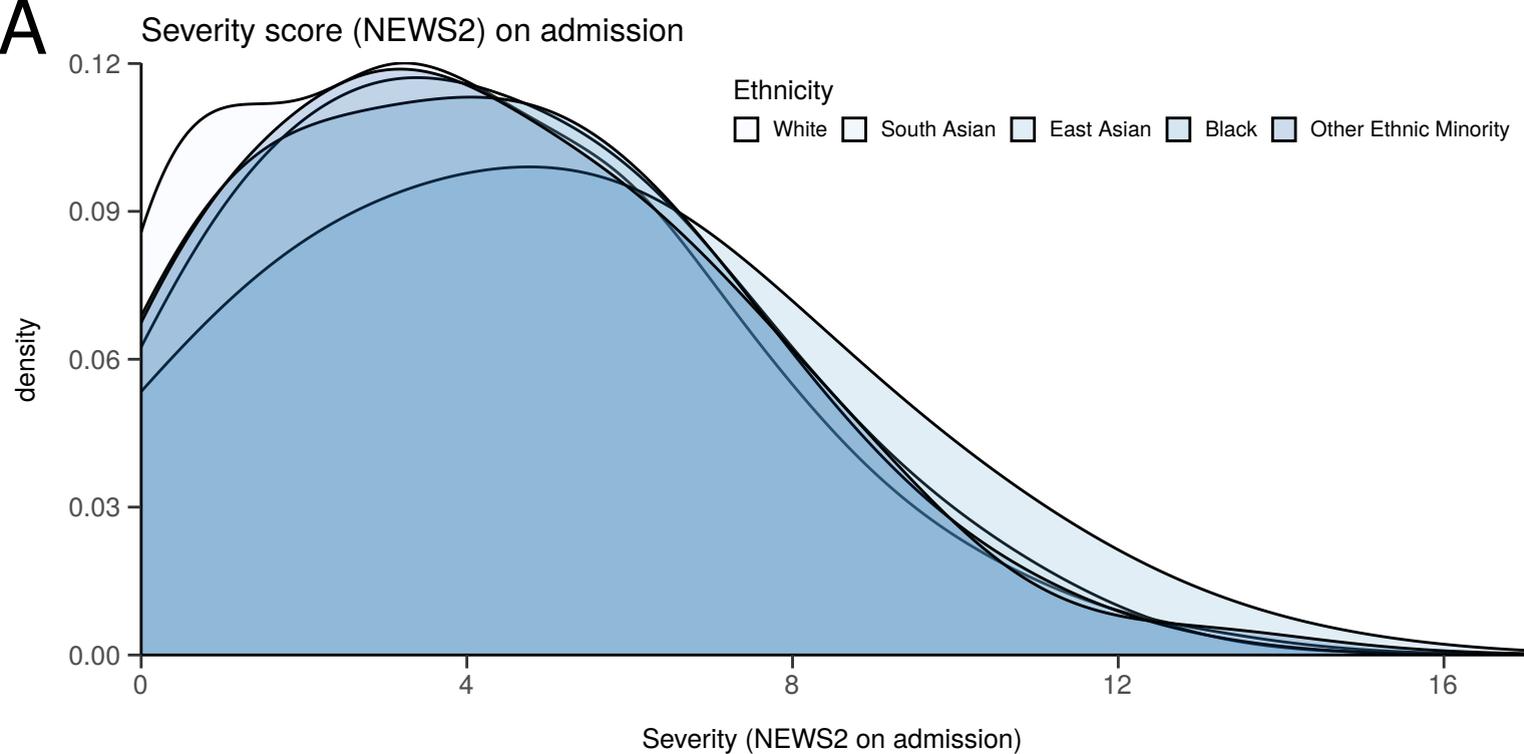
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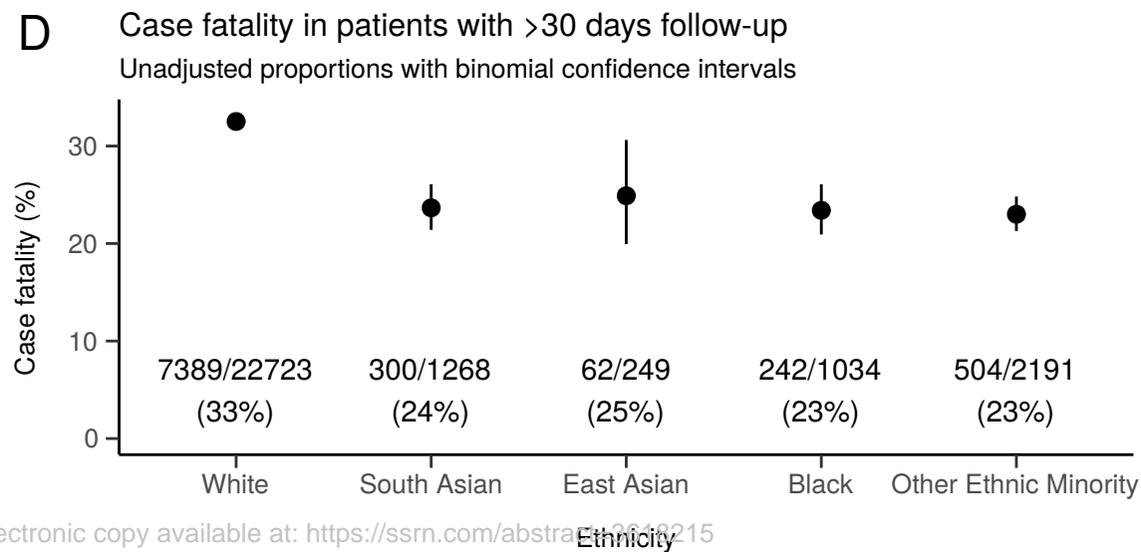
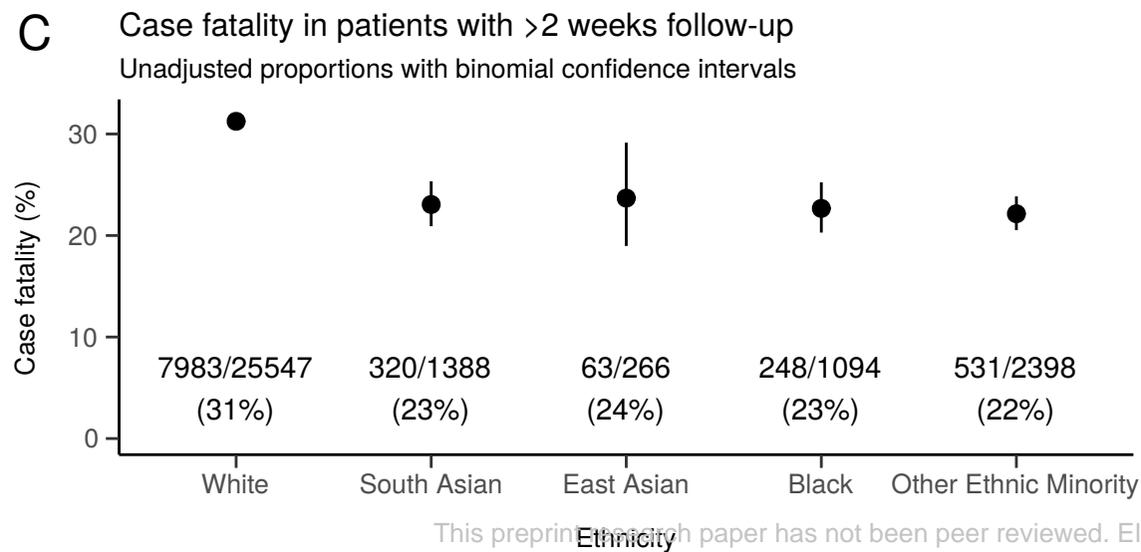
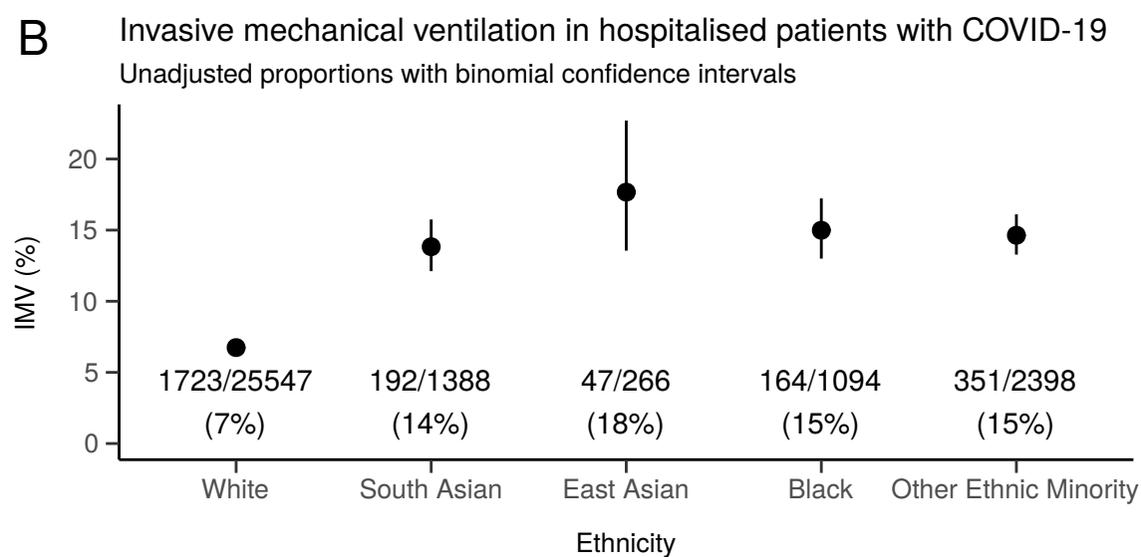
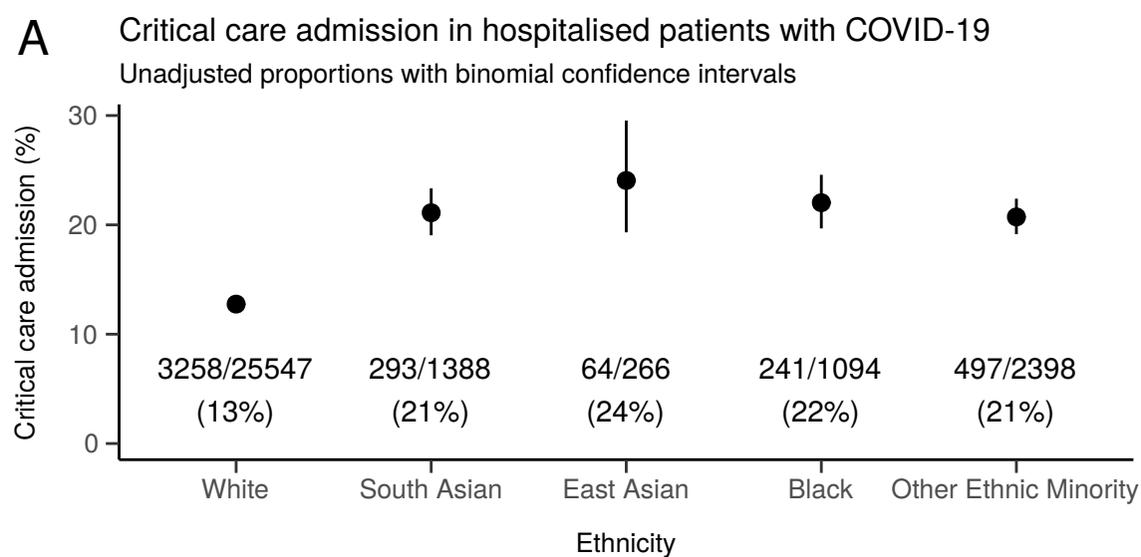
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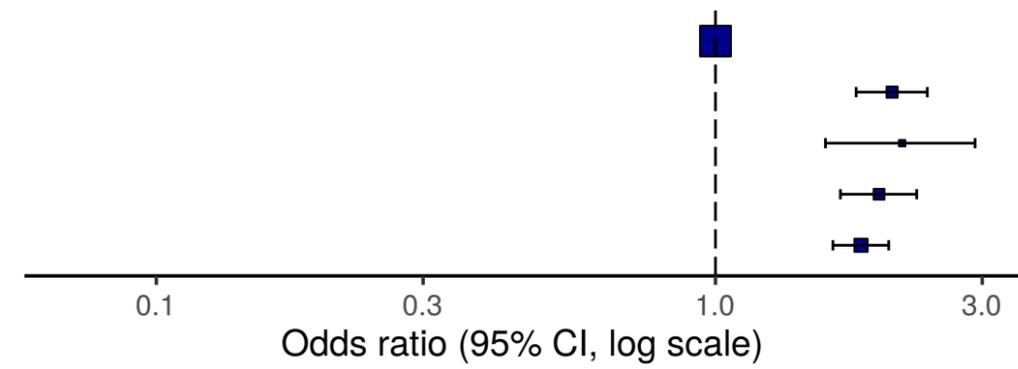






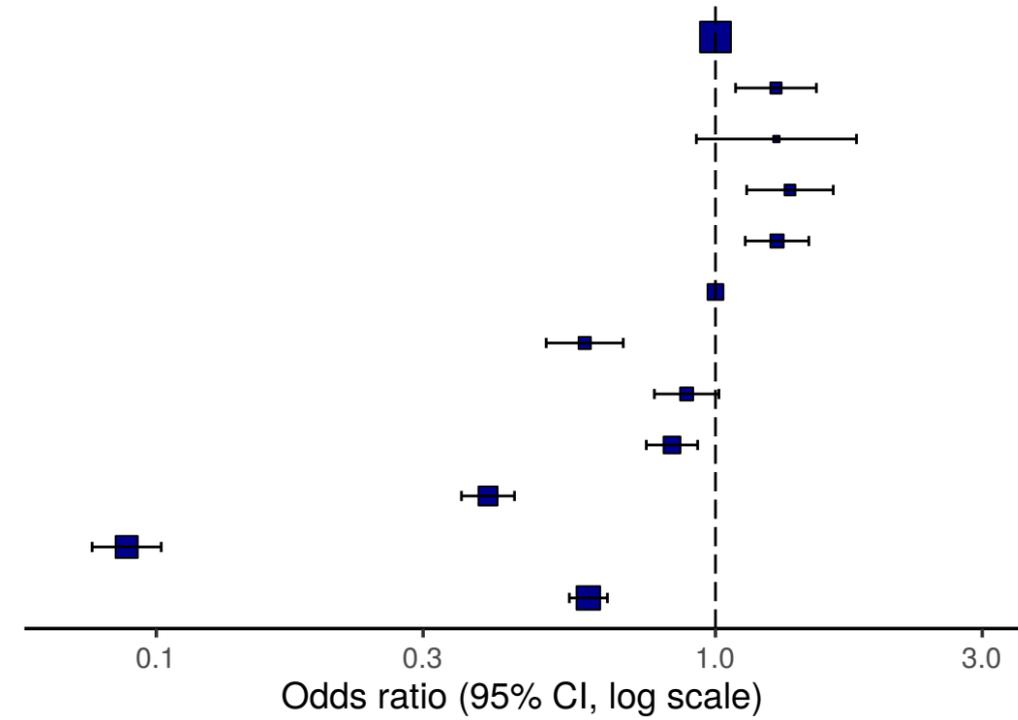
A Critical care admission: ethnicity alone

Ethnicity	White	-
	South Asian	2.07 (1.78-2.39, p<0.001)
	East Asian	2.16 (1.57-2.91, p<0.001)
	Black	1.96 (1.67-2.29, p<0.001)
	Other Ethnic Minority	1.82 (1.62-2.04, p<0.001)



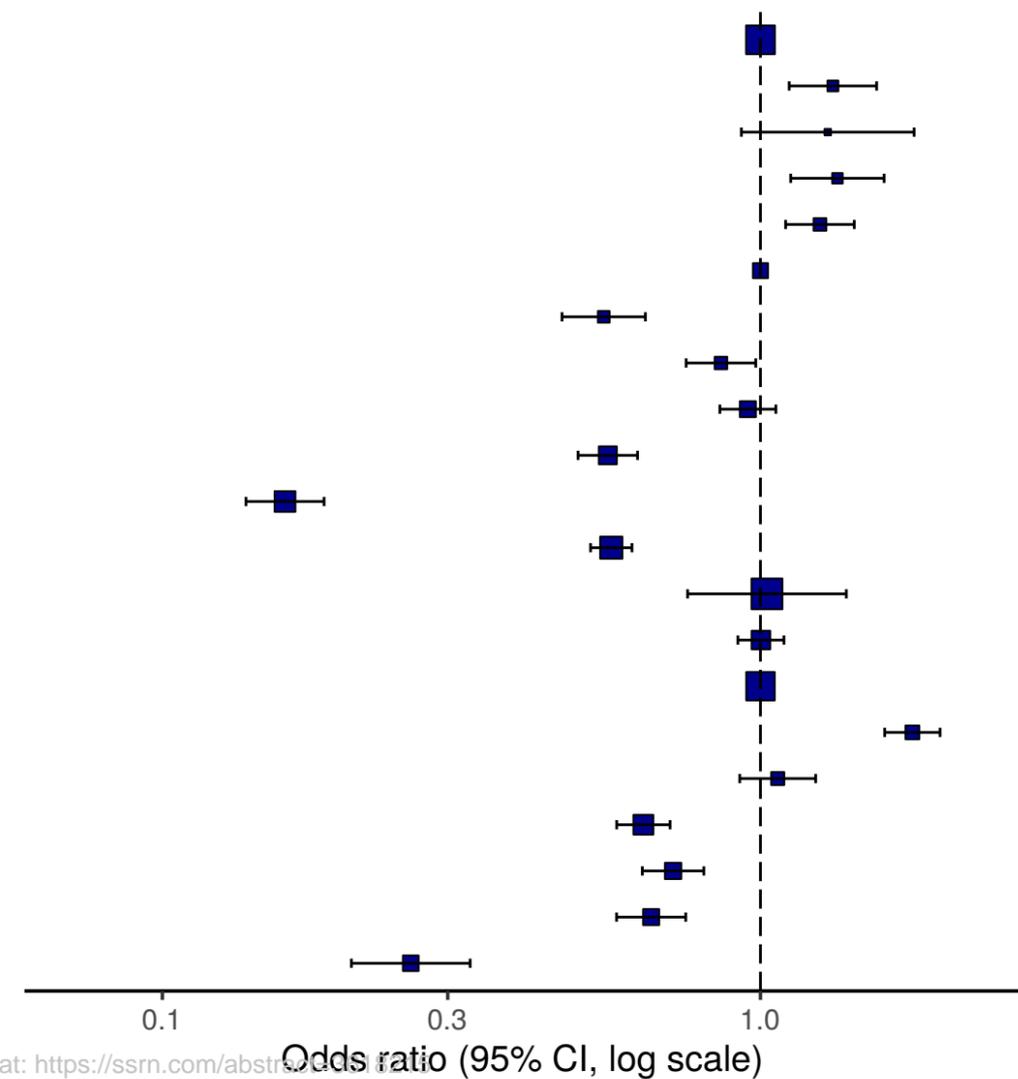
B Critical care: hierarchical baseline

Ethnicity	White	-
	South Asian	1.28 (1.09-1.52, p=0.003)
	East Asian	1.29 (0.92-1.79, p=0.136)
	Black	1.36 (1.14-1.62, p=0.001)
	Other Ethnic Minority	1.29 (1.13-1.47, p<0.001)
Age on admission (years)	50-59	-
	18-39	0.58 (0.50-0.68, p<0.001)
	40-49	0.89 (0.78-1.01, p=0.080)
	60-69	0.84 (0.75-0.93, p=0.001)
	70-79	0.39 (0.35-0.44, p<0.001)
	80+	0.09 (0.08-0.10, p<0.001)
Sex at Birth	Female	0.59 (0.55-0.64, p<0.001)



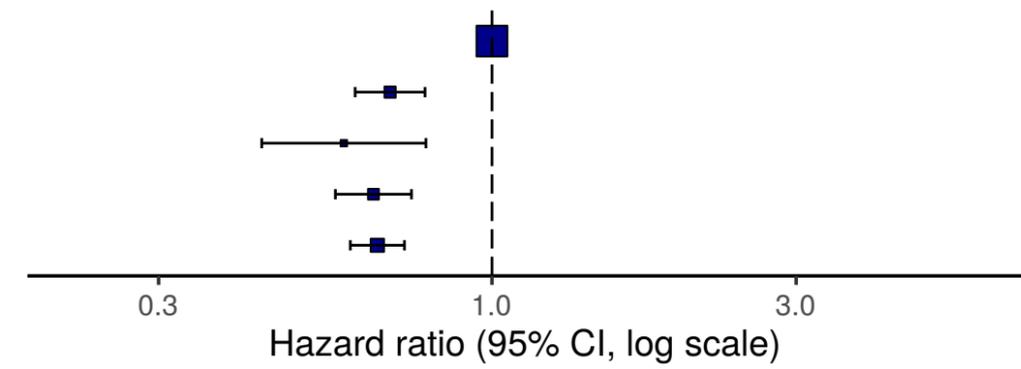
C Critical care: hierarchical with potential mediators

Ethnicity	White	-
	South Asian	1.32 (1.12-1.56, p=0.001)
	East Asian	1.30 (0.93-1.81, p=0.127)
	Black	1.34 (1.12-1.61, p=0.001)
	Other Ethnic Minority	1.26 (1.10-1.44, p=0.001)
Age on admission (years)	50-59	-
	18-39	0.55 (0.47-0.64, p<0.001)
	40-49	0.86 (0.75-0.98, p=0.026)
	60-69	0.95 (0.86-1.06, p=0.378)
	70-79	0.56 (0.50-0.62, p<0.001)
	80+	0.16 (0.14-0.19, p<0.001)
Sex at Birth	Female	0.56 (0.52-0.61, p<0.001)
Deprivation (IMD)	-	1.03 (0.76-1.39, p=0.873)
Diabetes	Yes	1.00 (0.92-1.09, p=0.969)
Obesity	No	-
	Yes	1.79 (1.61-2.00, p<0.001)
	(Missing)	1.07 (0.92-1.24, p=0.373)
Chronic cardiac disease	Yes	0.64 (0.57-0.71, p<0.001)
Chronic pulmonary disease	Yes	0.71 (0.63-0.80, p<0.001)
Chronic kidney disease	Yes	0.66 (0.57-0.75, p<0.001)
Dementia	Yes	0.26 (0.21-0.33, p<0.001)



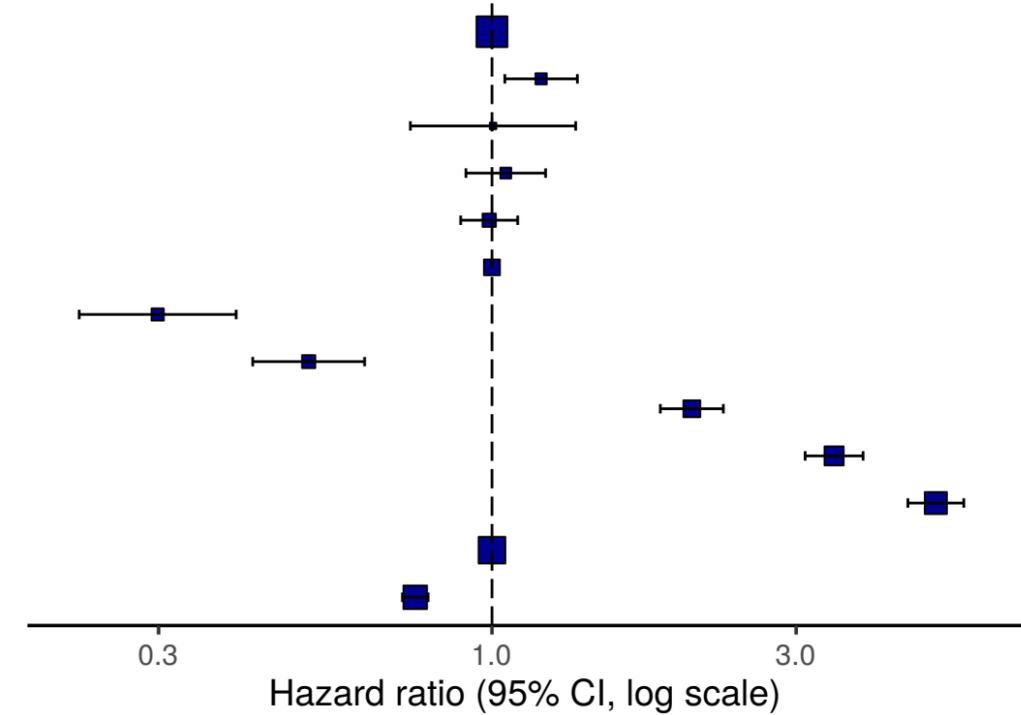
A In-patient survival: ethnicity alone

Ethnicity	White	-
	South Asian	0.69 (0.61-0.79, p<0.001)
	East Asian	0.59 (0.44-0.79, p<0.001)
	Black	0.65 (0.57-0.75, p<0.001)
	Other Ethnic Minority	0.66 (0.60-0.73, p<0.001)



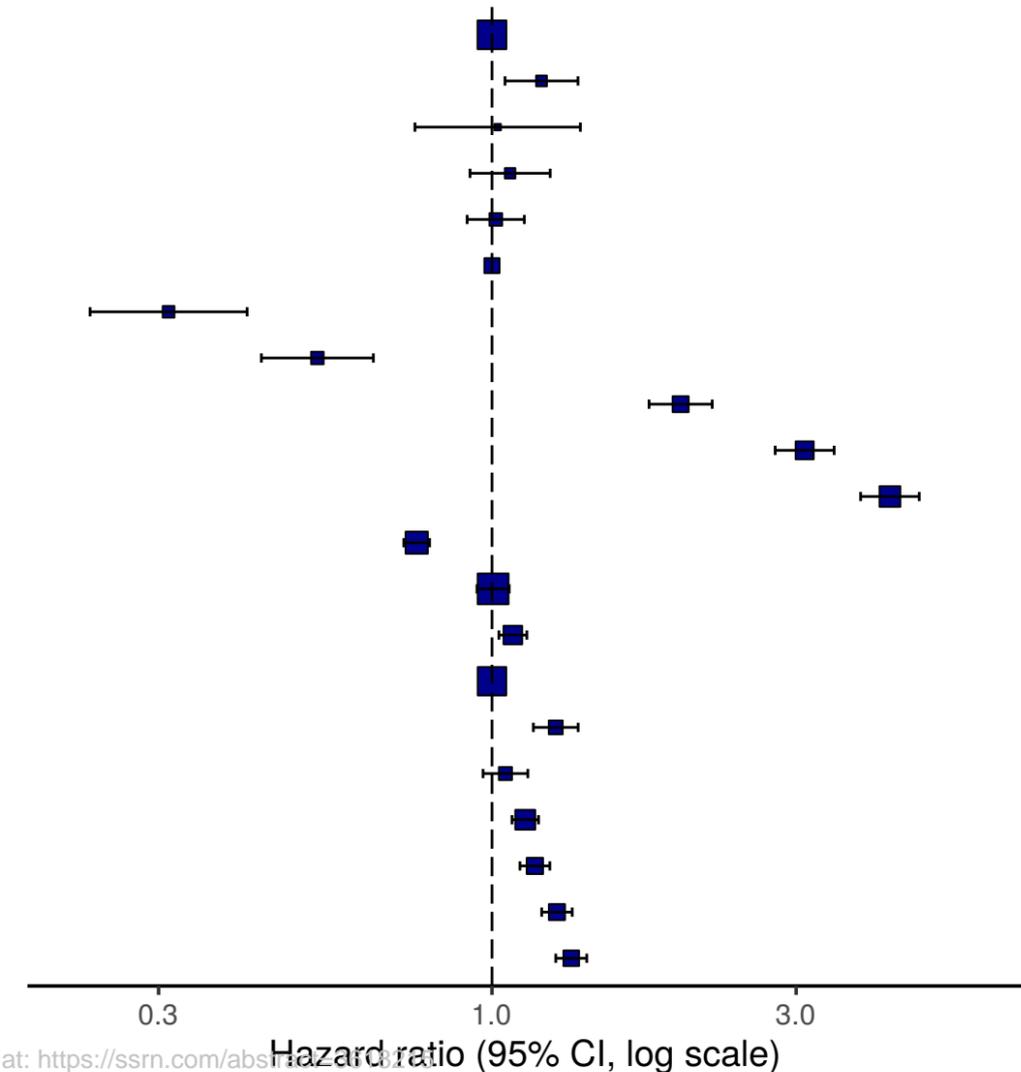
B In-patient survival: hierarchical baseline

Ethnicity	White	-
	South Asian	1.19 (1.05-1.36, p=0.008)
	East Asian	1.00 (0.74-1.35, p=0.980)
	Black	1.05 (0.91-1.21, p=0.500)
	Other Ethnic Minority	0.99 (0.89-1.10, p=0.850)
Age on admission (years)	50-59	-
	18-39	0.30 (0.23-0.40, p<0.001)
	40-49	0.52 (0.42-0.63, p<0.001)
	60-69	2.06 (1.84-2.31, p<0.001)
	70-79	3.44 (3.10-3.82, p<0.001)
	80+	4.97 (4.49-5.50, p<0.001)
Sex at Birth	Male	-
	Female	0.76 (0.72-0.79, p<0.001)

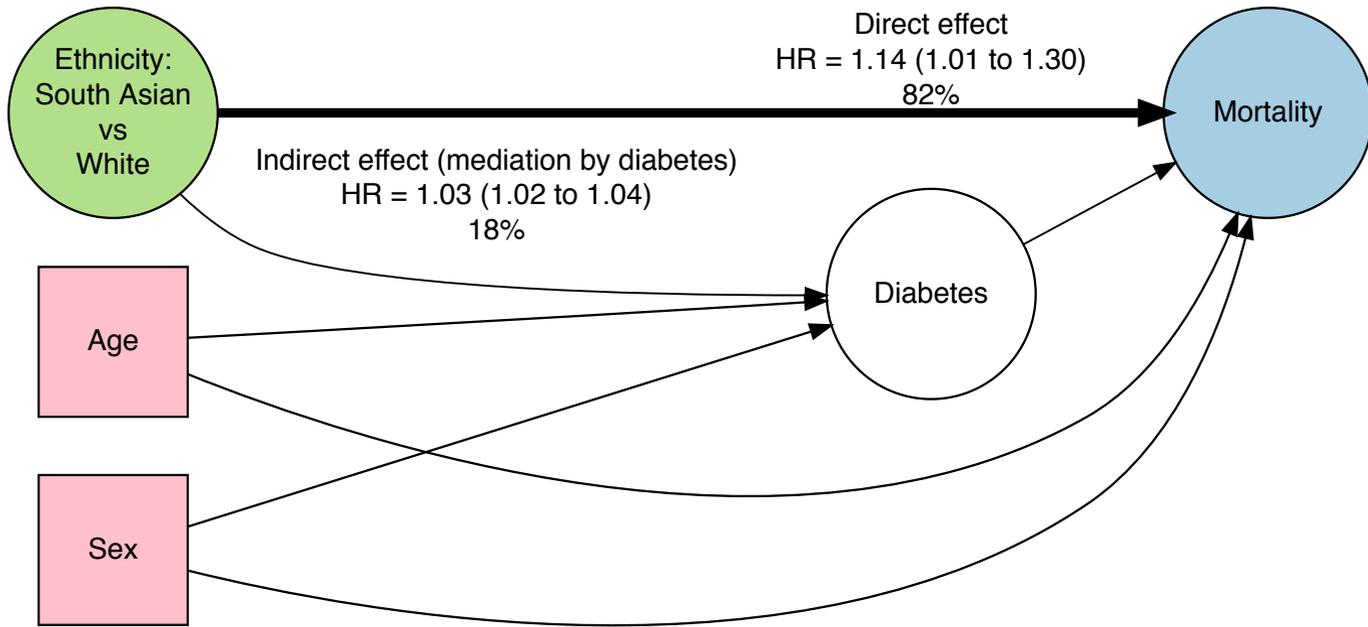


C In-patient survival: hierarchical with potential mediators

Ethnicity	White	-
	South Asian	1.20 (1.05-1.36, p=0.008)
	East Asian	1.02 (0.76-1.38, p=0.890)
	Black	1.07 (0.92-1.23, p=0.380)
	Other Ethnic Minority	1.01 (0.91-1.12, p=0.790)
Age on admission (years)	50-59	-
	18-39	0.31 (0.23-0.41, p<0.001)
	40-49	0.53 (0.43-0.65, p<0.001)
	60-69	1.98 (1.76-2.22, p<0.001)
	70-79	3.09 (2.78-3.44, p<0.001)
	80+	4.21 (3.79-4.68, p<0.001)
Sex at Birth	Female	0.76 (0.73-0.80, p<0.001)
Deprivation (IMD)	-	1.00 (0.95-1.06, p=0.890)
Diabetes	Yes	1.08 (1.03-1.13, p=0.003)
Obesity	No	-
	Yes	1.26 (1.16-1.37, p<0.001)
	(Missing)	1.05 (0.97-1.14, p=0.240)
Chronic cardiac disease	Yes	1.13 (1.08-1.18, p<0.001)
Chronic pulmonary disease	Yes	1.17 (1.11-1.23, p<0.001)
Chronic kidney disease	Yes	1.26 (1.20-1.34, p<0.001)
Dementia	Yes	1.33 (1.26-1.41, p<0.001)



A



B

